

CUMULATIVE DISTRIBUTION FUNCTION (C.D.F.):



IF X IS A RANDOM VARIABLE, LET

$$F(x) := P(X \leq x) \leftarrow \text{C.D.F. OF } X$$

PROPERTIES: $\lim_{x \rightarrow \infty} F(x) = 1$ $\lim_{x \rightarrow -\infty} F(x) = 0$

NON-DECREASING: $x_1 \leq x_2 \Rightarrow F(x_1) \leq F(x_2)$

RIGHT-CONTINUOUS: $\lim_{y \rightarrow x^+} F(y) = F(x)$ WHY?

THE EVENTS $\{X \leq y\}$ DECREASE MONOTONICALLY TO $\{X \leq x\}$ AS $y \searrow x$, THUS BY THE CONTINUITY OF MEASURE  

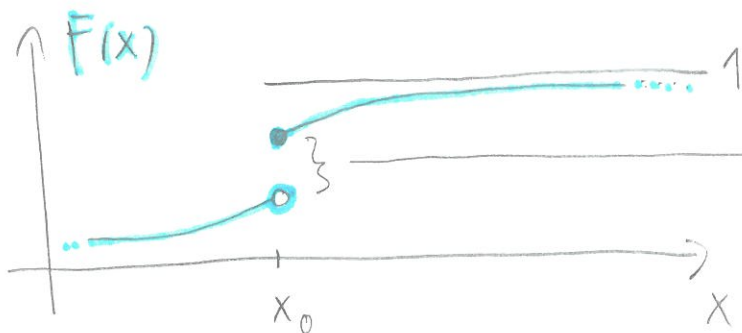
$$\lim_{y \searrow x} P(X \leq y) = P(X \leq x)$$

BUT: $\lim_{y \rightarrow x^-} F(y) = F(x_-) = P(X < x)$, THUS

$$P(X = x) = F(x) - F(x_-)$$
 THUS F HAS A

DISCONTINUITY AT x_0 IF AND ONLY IF

$$P(X = x_0) > 0. \quad \text{$$



$$P(X = x_0) = \text{size of jump}$$

NOTE: THERE ARE AT MOST COUNTABLY MANY VALUES OF x FOR WHICH $P(X = x) > 0$, BECAUSE FOR EACH SUCH x , THERE IS $q(x) \in \mathbb{Q}$ SUCH THAT $F(x-) < q(x) < F(x)$, AND \mathbb{Q} IS COUNTABLE.

DEF: (WEAK CONVERGENCE OF PROBABILITY DISTRIBUTIONS, CONV. IN DISTRIBUTION)

IF F_n IS THE C.D.F. OF X_n

F IS THE C.D.F. OF X

WE SAY THAT $F_n \Rightarrow F$, $X_n \Rightarrow X$ IF

$$\lim_{n \rightarrow \infty} F_n(x) = F(x)$$

FOR ALL $x \in \mathbb{R}$ FOR WHICH $P(X = x) = 0$

(F_n CONVERGES POINTWISE TO F AT THE POINTS OF CONTINUITY OF F)

NOTE: IF $P(X'_m = x_m) = 1$ (I.E.: X'_m IS DETERMINISTIC)

AND $P(X'_\infty = x_\infty) = 1$ THEN $F_m(x) = \begin{cases} 0 & \text{IF } x < x_m \\ 1 & \text{IF } x \geq x_m \end{cases}$

MOREOVER: $F_m \Rightarrow F_\infty \iff X_m \rightarrow X_\infty$. PROOF: EASY

NOTE: IF $X_m \rightarrow X_\infty$ THEN $F_m(x_\infty) = 0$, $F_\infty(x_\infty) = 1$,
 $F_m(x_\infty)$ DOES NOT CONVERGE TO $F_\infty(x_\infty)$ AS $m \rightarrow \infty$

CLAIM: IF $c \in \mathbb{R}$ THEN $X'_m \Rightarrow c$ IF AND

ONLY IF X'_m CONVERGES TO c IN PROBABILITY:

$\forall \varepsilon > 0$: $\lim_{n \rightarrow \infty} P(|X'_n - c| \geq \varepsilon) = 0$. PROOF: EASY

THUS WEAK LAW OF LARGE NUMBERS


CAN BE REPHRASED USING THE NOTION OF WEAK CONVERGENCE:

THM: IF X_1, X_2, \dots ARE I.I.D., $E(|X_1|) < +\infty$
 $E(X_1) = \mu$, $S_n = X_1 + \dots + X_n$, THEN

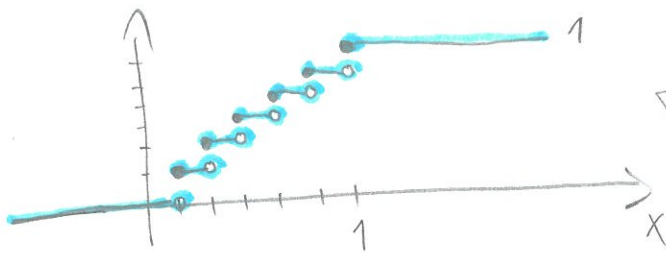
$$\frac{S_n}{n} \Rightarrow \mu$$

PROOF: IN A FEW WEEKS.



EX: IF X'_m IS UNIFORMLY DISTRIBUTED ON $\{1, 2, \dots, m\}$, I.E. $P(X'_m = k) = \frac{1}{m}$, $1 \leq k \leq m$

THEN $\frac{X'_m}{m} \Rightarrow \text{UNI}[0, 1]$ 

PROOF: $F_m(x) = P\left(\frac{X'_m}{m} \leq x\right) = \begin{cases} 0 & \text{IF } x \leq 0 \\ \frac{\lfloor m \cdot x \rfloor}{m} & \text{IF } 0 \leq x \leq 1 \\ 1 & \text{IF } x \geq 1 \end{cases}$



← GRAPH OF $F_m(x)$

$\forall x \in \mathbb{R} \quad \lim_{x \rightarrow \infty} F_m(x) = F(x) = \begin{cases} 0 & \text{IF } x \leq 0 \\ x & \text{IF } 0 \leq x \leq 1 \\ 1 & \text{IF } x \geq 1 \end{cases}$  

COUNTER-EXAMPLE:

IF WE CONSIDER X'_m FROM THE PREVIOUS EX., THEN $G_m(x) := P(X'_m \leq x) = \begin{cases} 0 & \text{IF } x \leq 0 \\ \frac{\lfloor x \rfloor}{m} & \text{IF } 0 \leq x \leq m \\ 1 & \text{IF } x \geq m \end{cases}$

THUS $\forall x \in \mathbb{R} \quad \lim_{m \rightarrow \infty} G_m(x) = 0$

BUT $G(x) \equiv 0$ IS NOT THE C.D.F. OF A R.V.,

THUS X'_m DOES NOT CONVERGE WEAKLY.

(MASS ESCAPES TO INFINITY)

EX: LET $X_{1,1}, X_{1,2}, \dots$ I.I.D. $\text{EXP}(1)$.

LET $M_m := \max \{ X_{1,1}, \dots, X_{1,m} \}$.

M_m GROWS TO INFINITY, BUT HOW FAST?

CLAIM: $M_m - \ln(m)$ CONVERGES IN DISTRIBUTION:

LET $F_m(x) := \mathbb{P}(M_m - \ln(m) \leq x) =$

$$= \mathbb{P}(M_m \leq \ln(m) + x) = \mathbb{P}\left(\bigcap_{i=1}^m \{X_{1,i} \leq \ln(m) + x\}\right)$$

$$= \prod_{i=1}^m \mathbb{P}(X_{1,i} \leq \ln(m) + x) = (1 - \exp(-(\ln(m) + x)))^m$$

$$= \left(1 - \frac{e^{-x}}{m}\right)^m, \text{ THUS } \lim_{m \rightarrow \infty} F_m(x) = \exp(-e^{-x})$$

$F(x) = \exp(-e^{-x})$ IS INDEED A C.D.F., IT IS

THE C.D.F. OF THE STANDARD GUMBEL
DISTRIBUTION.

EXTREME VALUE THEORY (BMETE 95MM16)

→ THEORY OF WEAK LIMITS OF THE MAXIMUM
OF RANDOM VARIABLES. USED IN
ACTUARIAL SCIENCE (INSURANCE MATH).

FUN FACT: IF Y_1, Y_2 I.I.D. STANDARD GUMBEL:

THEN $\max\{Y_1, Y_2\} \stackrel{A}{\sim} Y + \ln(2)$ WHERE Y IS

INDEED: $F^2(x) = F(x - \ln(2))$ ✓

ALSO: IF $n \gg 1$ THEN

$\max\{X_{1,1}, \dots, X_{2,n}\} = \max\{X_{1,1}, \dots, X_{n,1}\} \vee \max\{X_{n+1,1}, \dots, X_{2n,1}\}$
 $\approx Y + \ln(2n)$ $\approx Y_1 + \ln(n)$ $\approx Y_2 + \ln(n)$
 $\ln(2) + \ln(n)$

CLAIM: IF $X_{1,1}, X_{2,1}, \dots$ ARE \mathbb{Z} -VALUED R.V.'S

THEN $X_{1,n} \Rightarrow X_1 \iff \mathbb{P}(X_{1,n} = k) \xrightarrow{n \rightarrow \infty} \mathbb{P}(X_1 = k)$
 $k \in \mathbb{Z}$

PROOF: (A) \Rightarrow (B): $F_n(k - \frac{1}{2}) \rightarrow F(k - \frac{1}{2})$ AS $n \rightarrow \infty$ BY (A)

$$\mathbb{P}(X_{1,n} = k) = F_n((k+1) - \frac{1}{2}) - F_n(k - \frac{1}{2})$$

$$\mathbb{P}(X_1 = k) = F((k+1) - \frac{1}{2}) - F(k - \frac{1}{2})$$

THUS (B) HOLDS ✓

(B) \Rightarrow (A): NEXT PAGE

$$B \Rightarrow A: P(X_n \leq x) = \sum_{k \leq x} P(X_n = k)$$

FATOU

$$\liminf_{n \rightarrow \infty} P(X_n \leq x) = \liminf_{n \rightarrow \infty} \sum_{k \leq x} P(X_n = k) \geq$$

$$\sum_{k \leq x} P(X = k) = P(X \leq x), \text{ THUS}$$

$$\boxed{\liminf_{n \rightarrow \infty} P(X_n \leq x) \geq P(X \leq x)} \leftarrow \text{😊}$$

SIMILARLY:

$$\liminf_{n \rightarrow \infty} P(X_n > x) \geq P(X > x)$$

$$1 - P(X_n \leq x) \quad 1 - P(X \leq x), \text{ THUS}$$

$$\liminf_{n \rightarrow \infty} (-P(X_n \leq x)) \geq -P(X \leq x)$$

$$-\limsup_{n \rightarrow \infty} P(X_n \leq x), \text{ THUS}$$

$$\boxed{\limsup_{n \rightarrow \infty} P(X_n \leq x) \leq P(X \leq x)} \leftarrow \text{😞}$$

PUTTING 😊 AND 😞 TOGETHER:

$$\boxed{\lim_{n \rightarrow \infty} P(X_n \leq x) = P(X \leq x)}$$

THUS (A) HOLDS!