

PROOF OF Hoeffding USING LEMMA:

$$\begin{aligned} \mathbb{E}(e^{\lambda \cdot S_n}) &= \prod_{k=1}^n \mathbb{E}(e^{\lambda \cdot X_k}) \leq \prod_{i=1}^n \exp\left(\frac{1}{8} \cdot \lambda^2 \cdot (b_i - a_i)^2\right) = \\ &= e^{\frac{1}{2} \sigma^2 \cdot \lambda^2}, \text{ WHERE } \sigma^2 = \frac{1}{4} \cdot ((b_1 - a_1)^2 + \dots + (b_n - a_n)^2) \end{aligned}$$

$$\mathbb{P}(S_n \geq t) = \mathbb{E}(e^{\lambda \cdot S_n} \geq e^{\lambda t}) \leq \frac{\mathbb{E}(e^{\lambda S_n})}{e^{\lambda t}} \leq e^{\frac{1}{2} \sigma^2 \lambda^2 - \lambda t}$$

NOW  $\min_{\lambda \geq 0} \left\{ \frac{1}{2} \sigma^2 \cdot \lambda^2 - \lambda t \right\} = -\frac{1}{2} \frac{t^2}{\sigma^2}$  ✓

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EX: TOWN WITH 1000 HOUSEHOLDS, ONE GARBAGE CAN PER HOUSEHOLD. CAPACITY OF A GARBAGE CAN: 30 kg. THE AVERAGE WEEKLY GARBAGE OUTPUT MAY DIFFER FROM HOUSEHOLD TO HOUSEHOLD. AVERAGE WEEKLY OUTPUT OF TOWN IS  $10^4$  kg. CAPACITY OF ONE TRUCK:  $10^3$  kg

HOW MANY TRUCKS DO WE NEED IF WE WANT ALL GARBAGE TAKEN AWAY WITH 99% CHANCE?  
(GARBAGE TRUCK COMES ONCE PER WEEK)

SOLUTION:  $a_i = 0, b_i = 30$

$$\mathbb{P}(S_{1000} \geq 10^4 + t) \leq \exp\left(-\frac{2t^2}{1000 \cdot 30^2}\right) = 0.01 \Rightarrow$$

$$\Rightarrow t = \sqrt{\frac{1}{2} \ln(100) \cdot 9 \cdot 10^5} \approx 1440$$

WE NEED 12 TRUCKS.

THM: (BERNSTEIN'S INEQUALITY, 1937)

$X_{1,1}, X_{1,2}, \dots, X_{1,m}$  INDEPENDENT

$$S_m = X_{1,1} + \dots + X_{1,m}$$

$$D_m^2 := \text{Var}(S_m) = \text{Var}(X_{1,1}) + \dots + \text{Var}(X_{1,m})$$

IF  $P(|X_{1,r} - E(X_{1,r})| \leq K) = 1, 1 \leq r \leq m$ , THEN

$$P(S_m - E(S_m) \geq t \cdot D_m) \leq \exp\left(-\frac{t^2}{2 \cdot \left(1 + \frac{t \cdot K}{2D_m}\right)^2}\right)$$

FOR ANY  $0 \leq t \leq \frac{D_m}{K}$



COROLLARY: (BY APPLYING THM TO  $-S_m$ ):

$$P(S_m - E(S_m) \leq -t \cdot D_m) \leq \text{★}, \text{ THUS}$$

$$P(|S_m - E(S_m)| \geq t \cdot D_m) \leq 2 \cdot \text{★}$$

NOTE: BERNSTEIN IS QUITE SHARP IF

$X_{1,1}, X_{1,2}, \dots$  ARE I.I.D.,  $n \rightarrow \infty$ , BUT  $t$  IS FIXED:

$$\limsup_{n \rightarrow \infty} P\left(\frac{S_m - E(S_m)}{D_m} \geq t\right) \leq \exp\left(-\frac{t^2}{2}\right)$$

COMPARE THIS TO ...

COMPARE THIS TO CENTRAL LIMIT THM:

$$\lim_{n \rightarrow \infty} P\left(\frac{S_n - E(S_n)}{D_n} \geq t\right) = \int_t^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \leq$$

$$\leq \int_t^{\infty} e^{-x^2/2} dx \leq \int_t^{\infty} x \cdot e^{-x^2/2} dx = e^{-t^2/2}$$

$\uparrow$   
 $t \geq 1$

PROOF OF BERNSTEIN: W.L.O.G.:  $E(X_i) = 0$

LEMMA: IF  $P(|X| \leq K) = 1$ ,  $\text{Var}(X) = \sigma^2$ ,  $E(X) = 0$ ,

THEN

$$\hat{I}(\lambda) \leq \frac{\lambda^2 \cdot \sigma^2}{2} \cdot \left(1 + \frac{\lambda \cdot K \cdot e^{\lambda K}}{3}\right)$$

PROOF:  $Z(\lambda) = \sum_{r=0}^{\infty} \frac{\lambda^r}{r!} E(X^r) =$

$$= 1 + 0 + \frac{\lambda^2}{2} \cdot \sigma^2 + \sum_{r=3}^{\infty} \frac{\lambda^r}{r!} E(X^r) = \text{☆}$$

$k \geq 3$ :  $E(X^r) \leq K^{r-2} \cdot E(X^2) = K^{r-2} \cdot \sigma^2$ , THUS

$$\frac{\lambda^r}{r!} E(X^r) \leq \frac{\lambda^2 \cdot \sigma^2}{2} \cdot \frac{\lambda \cdot K}{3} \cdot \underbrace{\frac{(\lambda \cdot K)^{r-3}}{4 \cdot 5 \cdot \dots \cdot k}}_{\geq (r-3)!}$$

$$\text{☆} \leq 1 + \frac{\lambda^2}{2} \cdot \sigma^2 \cdot \left(1 + \frac{\lambda K}{3} \sum_{r=3}^{\infty} \frac{(\lambda \cdot K)^{r-3}}{(r-3)!}\right) = \text{☺}$$

$$\textcircled{\text{smiley}} = 1 + \frac{\lambda^2}{2} \cdot \sigma^2 \cdot \left(1 + \frac{\lambda K}{3} e^{\lambda K}\right) \leq \boxed{\text{SINCE } 1+x \leq e^x}$$

$$\leq \exp\left(\frac{\lambda^2}{2} \cdot \sigma^2 \cdot \left(1 + \frac{\lambda K}{3} e^{\lambda K}\right)\right) \checkmark$$

PROOF OF BERNSTEIN USING LEMMA:

$$\mathbb{E}\left(e^{\lambda \cdot S_n}\right) = \prod_{i=1}^n \mathbb{E}\left(e^{\lambda \cdot X_i}\right) \leq \exp\left(\frac{\lambda^2}{2} \cdot D_n \cdot \left(1 + \frac{\lambda K}{3} \cdot e^{\lambda K}\right)\right)$$

$$\mathbb{E}\left(e^{\lambda \cdot S_n / D_n}\right) \stackrel{\text{HW 1.3 (a)}}{\leq} \exp\left(\frac{\lambda^2}{2} \cdot \left(1 + \frac{\lambda \cdot K / D_n}{3} \cdot e^{\lambda \cdot K / D_n}\right)\right)$$

$$\mathbb{P}\left(S_n \geq t \cdot D_n\right) = \mathbb{P}\left(e^{\lambda \cdot S_n / D_n} > e^{\lambda t}\right) \stackrel{\text{MARKOV INEQ.}}{\leq}$$

$$\leq \inf_{\lambda > 0} \exp\left(\frac{\lambda^2}{2} \cdot \left(1 + \frac{\lambda \cdot K / D_n}{3} \cdot e^{\lambda K / D_n}\right) - \lambda \cdot t\right)$$

LET  $\lambda := \frac{t}{\left(1 + \frac{t \cdot K}{2 D_n}\right)^2}$

NOTE:  $\lambda < t$

ASSUMED:  $t \leq \frac{D_n}{K}$   $\Leftrightarrow \frac{t \cdot K}{D_n} \leq 1$

THEN:

$$\frac{\lambda^2}{2} \cdot \left(1 + \frac{\lambda \cdot K}{D_n} \cdot \frac{e^{\lambda K / D_n}}{3}\right) - \lambda \cdot t =$$

$$\frac{t^2}{2 \cdot \left(1 + \frac{t \cdot K}{2 D_n}\right)^2} \cdot \frac{1 + \frac{\lambda K}{D_n} \cdot \frac{e^{\lambda K / D_n}}{3}}{\left(1 + \frac{t \cdot K}{2 D_n}\right)^2} - \frac{t^2}{\left(1 + \frac{t \cdot K}{2 D_n}\right)^2} = \textcircled{\star}$$

$$\textcircled{\star} = \frac{t^2}{2 \cdot \left(1 + \frac{t \cdot K}{2D_m}\right)^2} \cdot \frac{1 + \frac{\lambda K}{D_m} \cdot \frac{e^{\lambda K/D_m}}{3}}{\left(1 + \frac{t \cdot K}{2D_m}\right)^2} - \frac{t^2}{\left(1 + \frac{t \cdot K}{2D_m}\right)^2}$$

NOTE:  $e^{\lambda K/D_m} \stackrel{\lambda < t}{\leq} e^{t \cdot K/D_m} \leq e^1 < 3$ , THUS

$$\frac{1 + \frac{\lambda K}{D_m} \cdot \frac{e^{\lambda K/D_m}}{3}}{\left(1 + \frac{t \cdot K}{2D_m}\right)^2} \leq \frac{1 + \frac{\lambda \cdot K}{D_m}}{1 + \frac{t \cdot K}{D_m} + \left(\frac{t \cdot K}{2D_m}\right)^2} \stackrel{\lambda < t}{\leq} 1$$

$$\textcircled{\star} \leq \frac{t^2}{2 \cdot \left(1 + \frac{t \cdot K}{2D_m}\right)^2} - \frac{t^2}{\left(1 + \frac{t \cdot K}{2D_m}\right)^2} = \frac{-t^2}{2 \cdot \left(1 + \frac{t \cdot K}{2D_m}\right)^2} \quad \checkmark$$

## SOME MEASURE THEORY:

FACT: □ MONOTONE CONVERGENCE THM:

□ IF  $0 \leq X_{1,n} \leq X_{1,2} \leq \dots \leq X_{1,m} \leq \dots$

THEN  $\mathbb{E} \left( \sup_n X_{1,n} \right) = \sup_n \mathbb{E} (X_{1,n})$

(AND EQUALITY HOLDS EVEN IN THE CASE WHEN  $\mathbb{E} \left( \sup_n X_{1,n} \right) = +\infty$ ) □

LEMMA (FATOU): IF  $X_n \geq 0 \quad \forall n \in \mathbb{N}$

THEN  $E \left( \liminf_{n \rightarrow \infty} X_n \right) \leq \liminf_{n \rightarrow \infty} E(X_n)$

PROOF:  $\liminf_{n \rightarrow \infty} X_n = \sup_n \inf_{m \geq n} X_m$

LET  $Y_n := \inf_{m \geq n} X_m$ , THEN

$$0 \leq Y_1 \leq Y_2 \leq \dots$$

NOTE:  $Y_n \leq X_n$

$$E \left( \liminf_{n \rightarrow \infty} X_n \right) = E \left( \sup_n Y_n \right) \stackrel{\text{MONOTONE CONV.}}{=} \sup_n E(Y_n)$$

$$= \sup_n E(Y_n) = \lim_{n \rightarrow \infty} E(Y_n) \leq \liminf_{n \rightarrow \infty} E(X_n)$$

THM (DOMINATED CONVERGENCE):

IF  $\lim_{n \rightarrow \infty} X_n$  EXISTS AND  $\forall n \in \mathbb{N}$

$|X_n| \leq Y$  WHERE  $E(Y) < +\infty$  THEN

$$E \left( \lim_{n \rightarrow \infty} X_n \right) = \lim_{n \rightarrow \infty} E(X_n)$$

PROOF: APPLY FATOU TO  $Y + X_n$  AND  $Y - X_n$

$$X := \lim_{n \rightarrow \infty} X_n$$

$$Z_n := Y + X_n$$

NOTE:  $Z_n \geq 0$

$$E(Y) + E(X) = E(Y + X) = E(\lim_{n \rightarrow \infty} Z_n) =$$

$$E(\liminf_{n \rightarrow \infty} Z_n) \stackrel{\text{FATOU}}{\leq} \liminf_{n \rightarrow \infty} E(Z_n) =$$

$$\liminf_{n \rightarrow \infty} (E(Y) + E(X_n)) = E(Y) + \liminf_{n \rightarrow \infty} E(X_n)$$

THUS  $E(X) \leq \liminf_{n \rightarrow \infty} E(X_n)$  ☺

NOW LET  $\tilde{Z}_n := Y - X_n$  NOTE:  $\tilde{Z}_n \geq 0$

$$E(Y) - E(X) = E(\liminf_{n \rightarrow \infty} \tilde{Z}_n) \stackrel{\text{FATOU}}{\leq}$$

$$\leq \liminf_{n \rightarrow \infty} E(\tilde{Z}_n) = \liminf_{n \rightarrow \infty} (E(Y) - E(X_n))$$

$$= E(Y) + \liminf_{n \rightarrow \infty} (-E(X_n)) = E(Y) - \limsup_{n \rightarrow \infty} E(X_n)$$

$$-E(X) \leq -\limsup_{n \rightarrow \infty} E(X_n) \Rightarrow E(X) \geq \limsup_{n \rightarrow \infty} E(X_n)$$

NOW ☹ AND ☺ TOGETHER GIVE: ☹

$$E(X) = \lim_{n \rightarrow \infty} E(X_n) \checkmark$$