# Noise and dynamical sensitivity in critical planar percolation

Based on C. Garban, G. Pete & O. Schramm: The Fourier spectrum of critical percolation, *Acta Math.* 2010.

### **Bernoulli**(p) **bond and site percolation**

Given an (infinite) graph G = (V, E) and  $p \in [0, 1]$ . Each site (or bond) is chosen open with probability p, closed with 1 - p, independently of each other. Consider the open connected clusters.  $\theta(p) := \mathbf{P}_p[0 \longleftrightarrow \infty]$ .



Theorem (Harris 1960 and Kesten 1980).  $p_c(\mathbb{Z}^2, \text{bond}) = p_c(\Delta, \text{site}) = 1/2$ , and  $\theta(1/2) = 0$ . For p > 1/2, there is a.s. one infinite cluster.

# **Crossing probabilities and criticality**



Theorem (Russo 1978 and Seymour-Welsh 1978). In critical percolation on almost any planar lattice, for L, n > 0,

 $0 < a_L < \mathbf{P}[$  left-right crossing in  $n \times Ln ] < b_L < 1.$ 

Same holds for annulus-crossings.

By repeating this on all scales, and gluing the pieces by FKG:  $(r/R)^{\alpha} < \mathbf{P}[\partial B_r \longleftrightarrow \partial B_R] < (r/R)^{\beta}.$  Moreover, for the (polychromatic)  $\ell$ -arm probabilities

$$\alpha_{\ell}(r,R) := \mathbf{P}[\partial B_r \xleftarrow{\ell} \partial B_R],$$

again have quasi-multiplicativity:  $\alpha_{\ell}(r,R) \simeq \alpha_{\ell}(r,\rho) \alpha_{\ell}(\rho,R)$ , and thus  $c_{\ell} (r/R)^{C_{\ell}} < \alpha_{\ell}(r,R) < C_{\ell} (r/R)^{c_{\ell}}$ .

But these are non-monotone events, so cannot use just FKG to prove this q-multiplicativity. Need Separation Lemma: conditioned on having  $\ell$  arms from  $\partial B_r$  to  $\partial B_R$ , the collection of interfaces both in  $B_R \setminus B_{R/2}$  and  $B_{2r} \setminus B_r$  are well-separated. Then we can glue.



# $\ensuremath{\mathsf{Bernoulli}}(1/2)$ bond and site percolation



## Conformal invariance on $\Delta$

**Theorem (Smirnov 2001).** For p = 1/2 site percolation on  $\Delta_{\eta}$ , and  $Q \subset \mathbb{C}$  a piecewise smooth quad (simply connected domain with four boundary points  $\{a, b, c, d\}$ ),

$$\lim_{\eta \to 0} \mathbf{P} \Big[ ab \longleftrightarrow cd \text{ inside } \mathcal{Q}, \text{ in percolation on } \Delta_{\eta} \Big]$$

exists, is strictly between 0 and 1, and conformally invariant.



Calls for a continuum scaling limit, encoding macroscopic connectivity, cluster boundaries, etc. Aizenman '95, Schramm '00, Camia-Newman '06, Sheffield '09, Schramm-Smirnov '10. In physics, correlation functions.

# $SLE_6$ exponents

Given the conformal invariance, the exploration path converges to the Stochastic Loewner Evolution with  $\kappa = 6$  (Schramm 2000).



Using the  $SLE_6$  curve, several critical exponents can be computed (Lawler-Schramm-Werner, Smirnov-Werner 2001, plus Kesten 1987), e.g.:

$$\alpha_4(r,R) := \mathbf{P}\left[\overbrace{r}^r\right] = (r/R)^{5/4 + o(1)},$$

 $\alpha_1(r,R) = (r/R)^{5/48+o(1)}, \text{ and } \theta(p_c+\epsilon) := \mathbf{P}_{p_c+\epsilon}[0 \longleftrightarrow \infty] = \epsilon^{5/36+o(1)}.$ Here  $\beta = 5/36 = \frac{5/48}{2-5/4} = \frac{\xi_1}{2-\xi_4}$ . (Will explain this numerology tomorrow.)

# **Percolation and noise**

Take an  $\omega$  critical percolation configuration. Let  $\omega^{\epsilon}$  be a new configuration, where each site (or bond) is resampled with probability  $\epsilon$ , independently. (The  $\epsilon$ -noised version of  $\omega$ .)



For how large an  $\epsilon$  can we still predict from  $\omega$  whether there is a left-right crossing in  $\omega^{\epsilon}$ ?

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For how large an  $\epsilon$  can we still predict from  $\omega$  whether there is a left-right crossing in  $\omega^{\epsilon}$ ?

### Naive idea: how many pivotals are there?

A site (or bond) is pivotal in  $\omega$ , if flipping it changes the existence of a left-right crossing.  $\mathbf{E}|\operatorname{Piv}_n| \simeq n^2 \alpha_4(n) \quad (= n^{3/4+o(1)}).$ 

Furthermore,  $\mathbf{E}[|\operatorname{Piv}_n|^2] \leq C (\mathbf{E}|\operatorname{Piv}_n|)^2$ . So,  $\mathbf{P}[|\operatorname{Piv}_n| > \lambda \mathbf{E}|\operatorname{Piv}_n|] < C/\lambda^2$ , any  $\lambda$ .

And not only  $\exists \epsilon \mathbf{P} [|\operatorname{Piv}_n| > \epsilon \mathbf{E} |\operatorname{Piv}_n|] > \epsilon$ , but  $\mathbf{P} [0 < |\operatorname{Piv}_n| < \epsilon \mathbf{E} |\operatorname{Piv}_n|] \simeq \epsilon^{11/9 + o(1)}$ , as  $\epsilon \to 0$  (exponent only for  $\Delta$ ).



Cannot have many pivotals  $\implies$  If  $\epsilon_n \mathbf{E}[|\operatorname{Piv}_n|] \rightarrow 0$ , then we don't hit any pivotals (even in expectation)  $\implies$  Asymptotically full correlation.

Cannot have few pivotals (if there's any)  $\implies$  If  $\epsilon_n \mathbf{E}[|\operatorname{Piv}_n|] \to \infty$ , hit many pivotals (at least in expectation). But  $\implies$  asymptotic independence!

### Noise sensitivity of percolation

All results use Fourier analysis of Boolean functions:

**Theorem (Benjamini, Kalai & Schramm 1998).** If  $\epsilon > 0$  is fixed, and  $f_n$  is the indicator function for a left-right percolation crossing in an  $n \times n$  square, then as  $n \to \infty$ 

$$\mathbf{E}[f_n(\omega)f_n(\omega^{\epsilon})] - \mathbf{E}[f_n(\omega)]^2 \to 0.$$

This holds for all  $\epsilon = \epsilon_n > c/\log n$ .

**Theorem (Schramm & Steif 2005).** Same if  $\epsilon_n > n^{-a}$  for some positive a > 0. If triangular lattice, may take any a < 1/8.

**Theorem (Garban, P & Schramm 2008).** Same holds if and only if  $\epsilon_n \mathbf{E}[|\text{pivotals}|] \to \infty$ . For triangular lattice, this threshold is  $\epsilon_n = n^{-3/4+o(1)}$ .

# **Dynamical percolation**

Each variable is resampled according to an independent Poisson(1) clock. This is a Markov process  $\{\omega(t) : t \in [0,\infty)\}$ , in which  $\omega(t+s)$  is an  $\epsilon$ -noised version of  $\omega(t)$ , with  $\epsilon = 1 - \exp(-s)$ .

An exceptional time is such a (random) t, at which an almost sure property of the static process fails for  $\omega(t)$ .

Main example: (Non-)existence of an infinite cluster in percolation.

**Toy example:** Brownian motion on the circle does sometimes hit a given point, as opposed to its static version: a uniform random point.

In this toy example, the set of exceptional times is a random Cantor set of Lebesgue measure zero (because of Fubini) and Hausdorff-dimension 1/2.

If the static event is not extremely unlikely, and it is very sensitive to noise, then we may have some chance to see an exceptional time.

### **Dynamical percolation results**

Theorem (Häggström, Peres & Steif 1997).

- No exceptional times when  $p \neq p_c$ .
- No exceptional times when  $p = p_c$  for bond percolation on  $\mathbb{Z}^d$ ,  $d \ge 19$ .

The latter is essentially due to Hara-Slade '90 on the off-critical exponent  $\beta = 1$ :



 $\theta(p_c + \epsilon) < C\epsilon$ , hence, even switching asymmetrically, **E**[number of  $\epsilon$ -subintervals of [0,1] with  $0 \leftrightarrow \infty$ ]  $\leq C$ . But this exceptional set is closed without isolated points, so this number should blow up, if non-zero.

#### Theorem (Schramm & Steif 2005).

- There are exceptional times (a.s.) for critical site percolation on the triangular lattice.
- They have Hausdorff dimension in [1/6, 31/36].

#### Theorem (Garban, P & Schramm 2008).

- There are exceptional times also on  $\mathbb{Z}^2$ .
- On the triangular grid they have Hausdorff dimension 31/36.
- On the triangular grid, there are exceptional times with an infinite white and an infinite black cluster simultaneously.  $(1/9 \le \dim \le 2/3)$

#### What is the Fourier spectrum and why is it useful?

 $f_n: \{\pm 1\}^{V_n} \longrightarrow \{\pm 1\}$  indicator of left-right crossing,  $V = V_n$  vertices.

 $(N_{\epsilon}f)(\omega) := \mathbf{E}[f(\omega^{\epsilon}) | \omega]$  is the noise operator, acting on the space  $L^{2}(\Omega, \mu)$ , where  $\Omega = \{\pm 1\}^{V}$ ,  $\mu$  uniform measure, inner product  $\mathbf{E}[fg]$ .

Correlation:  $\mathbf{E}[f(\omega^{\epsilon})f(\omega)] - \mathbf{E}[f(\omega)]\mathbf{E}[f(\omega^{\epsilon})] = \mathbf{E}[f(\omega)N_{\epsilon}f(\omega)] - \mathbf{E}[f(\omega)]^2$ . So, we would like to diagonalize the noise operator  $N_{\epsilon}$ .

Let  $\chi_i$  be the function  $\chi_i(\omega) = \omega(i)$ ,  $\omega \in \Omega$ .

For  $S \subset V$ , let  $\chi_S := \prod_{i \in S} \chi_i$ , the parity inside S. Then

$$N_{\epsilon}\chi_{i} = (1-\epsilon)\chi_{i}; \qquad N_{\epsilon}\chi_{S} = (1-\epsilon)^{|S|}\chi_{S}.$$

Moreover, the family  $\{\chi_S, S \subseteq V\}$  is an orthonormal basis of  $L^2(\Omega, \mu)$ .

Any function  $f \in L^2(\Omega, \mu)$  in this basis (Fourier-Walsh series):

$$\hat{f}(S) := \mathbf{E}[f\chi_S]; \qquad f = \sum_{S \subseteq V} \hat{f}(S) \chi_S.$$

The correlation:

$$\mathbf{E}[fN_{\epsilon}f] - \mathbf{E}[f]^{2} = \sum_{S} \sum_{S'} \hat{f}(S) \hat{f}(S') \mathbf{E}[\chi_{S} N_{\epsilon} \chi_{S'}] - \mathbf{E}[f\chi_{\emptyset}]^{2}$$
$$= \sum_{\emptyset \neq S \subseteq V} \hat{f}(S)^{2} (1-\epsilon)^{|S|} = \sum_{k=1}^{|V_{n}|} (1-\epsilon)^{k} \sum_{|S|=k} \hat{f}(S)^{2}.$$

By Parseval,  $\sum_{S} \hat{f}(S)^2 = \mathbf{E}[f^2] = 1$ . So can define probability measure  $\mathbf{P}[\mathscr{S}_f = S] := \hat{f}(S)^2 / \mathbf{E}[f^2]$ , the spectral sample  $\mathscr{S}_f \subseteq V$ .

If, for some functions  $f_n$  and numbers  $k_n$ , we have  $\mathbf{P}[0 < |\mathscr{S}_n| < tk_n] \to 0$ as  $t \to 0$ , uniformly in n, then  $(1 - \epsilon)^k \sim \exp(-\epsilon k)$  implies that for  $\epsilon_n \gg 1/k_n$  we have asymptotic independence. Maybe with  $k_n = \mathbf{E}|\mathscr{S}_n|$ ?

### **Pivotals versus spectral sample**

$$\begin{split} \nabla_{i}f(\omega) &:= f(\sigma_{i}(\omega)) - f(\omega) \in \{-2, 0, +2\} \text{ gradient.} \\ \nabla_{i}f(\omega) &= \sum_{S}\hat{f}(S)[\chi_{S}(\sigma_{i}(\omega)) - \chi_{S}(\omega)], \text{ hence } \widehat{\nabla_{i}f}(S) = -2\hat{f}(S)\mathbf{1}_{i\in S}. \\ \mathbf{P}[i \in \operatorname{Piv}_{f}] &= \frac{1}{4} \|\nabla_{i}f\|_{2}^{2} = \frac{1}{4}\sum_{S}\widehat{\nabla_{i}f}(S)^{2} = \sum_{S\ni i}\hat{f}(S)^{2} = \mathbf{P}[i \in \mathscr{S}_{f}]. \\ \text{Thus, } \mathbf{E}|\mathscr{S}_{f}| &= \mathbf{E}|\operatorname{Piv}_{f}|. \text{ So, the pivotal upper bound for noise sensitivity} \\ \text{ is sharp if there is tightness around } \mathbf{E}|\mathscr{S}|. \end{split}$$

Will see  $\mathbf{P}[i, j \in \operatorname{Piv}_f] = \mathbf{P}[i, j \in \mathscr{S}_f]$ , hence  $\mathbf{E}|\mathscr{S}_f|^2 = \mathbf{E}|\operatorname{Piv}_f|^2$ .

Not for more points and higher moments! Both random subsets measure the "influence" or "relevance" of bits, but in different ways.

For percolation,  $\mathbf{E}[|\operatorname{Piv}_n|^2] \leq C(\mathbf{E}|\operatorname{Piv}_n|)^2$ , hence  $\exists c > 0$  s.t.  $\mathbf{P}[|\mathscr{S}_n| > c \mathbf{E}|\mathscr{S}_n|] > c$ . That's why one hopes for tightness around mean.



# Three very simple examples

$$\begin{aligned} \mathsf{Dictator}_n(x_1, \dots, x_n) &:= x_1 \, . \\ \mathsf{Here } \mathsf{Cov} \big[ \mathsf{Dic}_n(x), \mathsf{Dic}_n(x^{\epsilon}) \big] &= 1 - \epsilon, \text{ so noise-stable.} \\ \mathsf{And } \mathbf{P} \big[ \mathscr{S}_n &= \{x_1\} \big] &= 1. \end{aligned}$$

$$\begin{aligned} \mathsf{Majority}_n(x_1, \dots, x_n) &:= \mathrm{sgn} \left( x_1 + \dots + x_n \right) \approx \frac{1}{\sqrt{n}} (x_1 + \dots + x_n) \,. \\ \mathsf{Here } \operatorname{Cov} \left[ \operatorname{Maj}_n(x), \operatorname{Maj}_n(x^{\epsilon}) \right] &= 1 - O(\epsilon), \text{ so noise-stable.} \\ \mathsf{And } \mathbf{P} \left[ \mathscr{S}_n &= \{x_i\} \right] &\asymp 1/n, \text{ most of the weight is on singletons.} \\ \mathsf{On the other hand, } \mathbf{E} |\mathscr{S}_n| &= \mathbf{E} |\operatorname{Piv}_n| \asymp \frac{1}{\sqrt{n}} n \asymp \sqrt{n}. \end{aligned}$$

 $\begin{aligned} &\mathsf{Parity}_n(x_1, \dots, x_n) := x_1 \cdots x_n \\ &\mathsf{Here } \mathsf{Cov} \big[ \mathsf{Par}_n(x), \mathsf{Par}_n(x^{\epsilon}) \big] = (1 - \epsilon)^n \text{, the most sensitive to noise.} \\ &\mathsf{And } \mathbf{P} \big[ \mathscr{S}_n = \{x_1, \dots, x_n\} \big] = 1. \end{aligned}$ 

#### Benjamini, Kalai & Schramm 1998

**Theorem.** A sequence  $f_n$  of monotone Boolean functions is noise sensitive, i.e., for any fixed  $\epsilon > 0$ ,

$$\mathbf{E}\left[f_n(\omega)f_n(\omega^{\epsilon})\right] - \mathbf{E}\left[f_n(\omega)\right]^2 \to 0$$

as  $n \to \infty$ , iff it is asymptotically uncorrelated with all weighted majorities  $\operatorname{Maj}_{w}(x_1, \ldots, x_n) = \operatorname{sign} \sum_{i=1}^{n} x_i w_i$ . Also, not very slow decorrelation with all subset-majorities is enough for sensitivity.

**Theorem.** The left-right percolation crossing in an  $n \times n$  square is noise sensitive, even with  $\epsilon = \epsilon_n > c/\log n$ .

# Schramm & Steif 2005

**Theorem.** If  $f : \Omega \longrightarrow \mathbb{R}$  can be computed with a randomized algorithm with revealment  $\delta$  (each bit is read only with probability  $\leq \delta$ ), then

 $\sum_{S:|S|=k} \hat{f}(S)^2 \leqslant \delta k \, \|f\|_2^2 \, .$ 

For left-right crossing in  $n \times n$  box on the hexagonal lattice, exploration interface with random starting point gives revealment  $n^{-1/4+o(1)}$  (it has length  $n^{7/4+o(1)}$ , given by 2-arm exponent), while  $\sum_{k \leq m} k \asymp m^2$ , thus:

**Theorem.** Left-right crossing on the triangular lattice is noise sensitive under  $\epsilon_n > n^{-a}$ , with any a < 1/8. Even on square lattice, can take some positive a > 0.

The revealment is at least  $n^{-1/2+o(1)}$  for any algorithm computing the crossing, hence this method can give only  $n^{-1/4+o(1)}$ -sensitivity, far from the conjectured  $\epsilon_n = n^{-3/4+o(1)}$ .

# The GPS approach, 2008

Although goal is to understand size, Gil Kalai suggested trying to understand entire distribution of  $\mathscr{S}_f$ . A strange random set of bits.

Effective sampling? If f is an effectively computable Boolean function, then there is an effective quantum algorithm for  $\mathscr{S}_f$  [Bernstein-Vazirani 1993].

For  $\mathscr{S}_{Q,n}$  (left-right crossing in a conformal rectangle Q, mesh 1/n), [Smirnov '01] + [Tsirelson '04] + [Schramm-Smirnov '11] implies that it has a conformally invariant scaling limit.

How to prove tightness for the size of strange random fractal-like sets?

#### Basic properties of the spectral sample

For  $A \subseteq V$ :  $\mathbf{E}[\chi_S \mid \mathscr{F}_A] = \begin{cases} \chi_S & S \subseteq A, \\ 0 & \text{otherwise}. \end{cases}$ Therefore,  $\mathbf{E}[f \mid \mathscr{F}_A] = \sum_{S \subseteq A} \widehat{f}(S) \chi_S$ , a nice projection. Also, for  $T \subseteq A$ :  $\mathbf{E}[f \chi_T \mid \mathscr{F}_{A^c}] = \sum_{S \subseteq A^c} \widehat{f}(T \cup S) \chi_S$ , hence  $\mathbf{E}[\mathbf{E}[f \chi_T \mid \mathscr{F}_{A^c}]^2] = \sum_{S \subseteq A^c} \widehat{f}(T \cup S)^2 = \mathbf{P}[\mathscr{S} \cap A = T].$ 

This is the Random Restriction Lemma of Linial-Mansour-Nisan '93. E.g.,

$$\begin{split} \mathbf{P}\big[i, j \in \mathscr{S}_{f}\big] &= \mathbf{E}\Big[\mathbf{E}\big[f\chi_{\{i, j\}} \mid \mathscr{F}_{\{i, j\}^{c}}\big]^{2}\Big] \\ &= \frac{1}{4}\mathbf{P}\big[\omega\big|_{\{i, j\}^{c}} \text{ is such that } i, j \text{ each may be pivotal}\,\big] \\ &= \mathbf{P}\big[i, j \in \operatorname{Piv}_{f}\big]\,. \end{split}$$

How does  $[\mathscr{S}_n \cap B \mid \mathscr{S}_n \cap B \neq \emptyset]$  look like?

 ${\cal B}$  as set has to be pivotal.



**Strong Separation Lemma.** For  $d(B, \partial Q) > diam(B)$ , conditioned on the 4 interfaces to reach  $\partial B$ , with *arbitrary starting points*, with a uniformly positive conditional probability the interfaces are well-separated around  $\partial B$ . Very bad separation is very unlikely. [Simple proof by Damron-Sapozhnikov '09, following Kesten '87. Also explained in Appendix to GPS '11.]

**Corollary 1.**  $\mathbf{P}\left[\mathscr{S}_n \cap B_r \neq \emptyset\right] \asymp \alpha_4(r, n)$ .

Corollary 2.  $\mathbf{E}\left[\left|\mathscr{S}_{n} \cap B_{r}\right| \mid \mathscr{S}_{n} \cap B_{r} \neq \emptyset\right] \asymp r^{2} \alpha_{4}(1,r) \asymp \mathbf{E}|\mathscr{S}_{r}|.$ 

# Self-similarity for left-right crossing of $n \times n$ square

$$\mathbf{E}|\mathscr{S}_{n}| = \mathbf{E}|\operatorname{Piv}_{n}| \asymp n^{2} \alpha_{4}(1, n) \stackrel{\Delta}{\asymp} n^{3/4+o(1)},$$
$$\mathbf{E}|\mathscr{S}_{n}(r)| := \mathbf{E}\left[\#\left\{r\text{-boxes } \mathscr{S}_{n} \cap B_{r} \neq \emptyset\right\}\right] \asymp \frac{n^{2}}{r^{2}} \alpha_{4}(r, n) \asymp \mathbf{E}|\mathscr{S}_{n/r}|,$$
$$\mathbf{E}\left[|\mathscr{S}_{n} \cap B_{r}| \mid \mathscr{S}_{n} \cap B_{r} \neq \emptyset\right] \asymp r^{2} \alpha_{4}(1, r) \asymp \mathbf{E}|\mathscr{S}_{r}|.$$

Of course,  $r^2 \alpha_4(1,r) \cdot \frac{n^2}{r^2} \alpha_4(r,n) \asymp n^2 \alpha_4(1,n)$ , by quasi-multiplicativity.

#### Self-similarity for left-right crossing of $n \times n$ square

$$\begin{split} \mathbf{E}|\mathscr{S}_{n}| &= \mathbf{E}|\operatorname{Piv}_{n}| \asymp n^{2} \alpha_{4}(1,n) \stackrel{\Delta}{\asymp} n^{3/4+o(1)}, \\ \mathbf{E}|\mathscr{S}_{n}(r)| &:= \mathbf{E}\Big[ \# \big\{ r\text{-boxes } \mathscr{S}_{n} \cap B_{r} \neq \emptyset \big\} \Big] \asymp \frac{n^{2}}{r^{2}} \alpha_{4}(r,n) \asymp \mathbf{E}|\mathscr{S}_{n/r}|, \\ \mathbf{E}\Big[ |\mathscr{S}_{n} \cap B_{r}| \ \Big| \ \mathscr{S}_{n} \cap B_{r} \neq \emptyset \Big] \asymp r^{2} \alpha_{4}(1,r) \asymp \mathbf{E}|\mathscr{S}_{r}|. \end{split}$$
Of course,  $r^{2} \alpha_{4}(1,r) \cdot \frac{n^{2}}{r^{2}} \alpha_{4}(r,n) \asymp n^{2} \alpha_{4}(1,n), \text{ by quasi-multiplicativity.} \end{split}$ 

Similar to the zero-set of simple random walk:  $\mathbf{E}|\mathcal{Z}_n| \asymp n n^{-1/2} = n^{1/2}$ ,

$$\mathbf{E}|\mathcal{Z}_{n}(r)| := \mathbf{E}\Big[\#\big\{r\text{-intervals } \mathcal{Z}_{n} \cap I_{r} \neq \emptyset\big\}\Big] \asymp \frac{n}{r} (n/r)^{-1/2} \asymp \mathbf{E}|\mathcal{Z}_{n/r}|,$$
$$\mathbf{E}\Big[|\mathcal{Z}_{n} \cap I_{r}| \mid \mathcal{Z}_{n} \cap I_{r} \neq \emptyset\Big] \asymp r r^{-1/2} \asymp \mathbf{E}|\mathcal{Z}_{r}|.$$

These results are related to the existence of scaling limits.

#### What concentration can we expect?

 $\mathscr{S}_n$  is very different from uniform set of similar density: i.i.d.  $\mathbf{P}[x \in \mathscr{U}_n] = n^{-5/4}$ . Hence  $\mathbf{E}|\mathscr{U}_n| = n^{3/4}$ .

For large  $r \gg n^{5/8}$ , this  $\mathscr{U}_n$  intersects every r-box; for small r, if it intersects one, there is just one point there.

Concentration of size: roughly within  $\sqrt{\mathbf{E}|\mathscr{U}_n|} = n^{3/8}$ .

A bit more similar: for  $i = 1, \ldots, (n/r)^2$ , i.i.d.  $\mathbf{P}[X_i = r^{3/4}] = (n/r)^{-5/4}$ ,  $X_i = 0$  otherwise. Then  $S_{n,r} := \sum_i X_i$ . Hence  $\mathbf{E}|S_{n,r}| = n^{3/4}$ .

For  $r = n^{\gamma}$ , size  $|S_{n,r}|$  is concentrated within  $n^{3/8(1+\gamma)}$ , still  $o(\mathbf{E}|S_{n,r}|)$ .

For self-similar sets, we expect only tightness around the mean:  $\mathbf{P}[0 < |\mathscr{S}_n| < \lambda \mathbf{E}|\mathscr{S}_n|] \to 0$  as  $\lambda \to 0$ , uniformly in n.

#### **Proving tightness with a lot of independence**

Assume we have the following ingredients, true for the zeroes:

(1) 
$$\mathbf{P}\Big[|\mathcal{Z}_n \cap I_r| > c \, \mathbf{E}|\mathcal{Z}_r| \ \Big| \ \mathcal{Z}_n \cap I_r \neq \emptyset, \ \mathscr{F}_{[n]\setminus I_r}\Big] \geqslant c > 0.$$

(2)  $\mathbf{P}[|\mathcal{Z}_n(r)| = k] \leq g(k) \mathbf{P}[|\mathcal{Z}_n(r)| = 1]$ , with sub-exponential g(k):

when the r-intervals intersected are scattered, have to pay k times to get to and leave them, and this cost is not balanced by combinatorial entropy.

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when the r-intervals intersected are scattered, have to pay k times to get to and leave them, and this cost is not balanced by combinatorial entropy.

$$\begin{split} \mathbf{P}\big[ \, 0 < |\mathcal{Z}_n| < c \, \mathbf{E}|\mathcal{Z}_r| \, \big] &= \sum_{k \ge 1} \mathbf{P}\Big[ \, 0 < |\mathcal{Z}_n| < c \, \mathbf{E}|\mathcal{Z}_r| \, , \ |\mathcal{Z}_n(r)| = k \, \Big] \\ & \text{by (1):} \quad \leqslant \sum_{k \ge 1} (1-c)^k \, \mathbf{P}\big[ \, |\mathcal{Z}_n(r)| = k \, \big] \\ & \text{by (2):} \quad \leqslant O(1) \, \mathbf{P}\big[ \, |\mathcal{Z}_n(r)| = 1 \, \big] \asymp (n/r)^{1-3/2}, \end{split}$$

which, using  $\lambda = \frac{c \mathbf{E}|\mathcal{Z}_r|}{\mathbf{E}|\mathcal{Z}_n|} \asymp (r/n)^{1/2}$ , reads as  $\mathbf{P} \left[ 0 < |\mathcal{Z}_n| < \lambda \mathbf{E}|\mathcal{Z}_n| \right] \asymp \lambda$ .

### But we know much less independence for $\mathscr{S}_n$

(1')  $\mathbf{P}\Big[|\mathscr{S}_n \cap B_r/3| > c \mathbf{E}|\mathscr{S}_r| \ \Big| \ \mathscr{S}_n \cap B_r \neq \emptyset = \mathscr{S}_n \cap W\Big] \ge c > 0,$ for any W that is not too close to  $B_r$ .

Why only this negative conditioning? Inclusion formula:

$$\mathbf{P}\big[\mathscr{S}_f \subset U\big] = \sum_{S \subset U} \hat{f}(S)^2 = \mathbf{E}\Big[\left(\sum_{S \subset U} \hat{f}(S) \,\chi_S\right)^2\Big] = \mathbf{E}\Big[\mathbf{E}\big[f \mid \mathscr{F}_U\big]^2\Big]$$

From this, for disjoint subsets A and B,

$$\mathbf{P}\left[\mathscr{S}_{f} \cap B \neq \emptyset = \mathscr{S}_{f} \cap A\right] = \mathbf{P}\left[\mathscr{S}_{f} \subseteq A^{c}\right] - \mathbf{P}\left[\mathscr{S}_{f} \subseteq (A \cup B)^{c}\right]$$
$$= \mathbf{E}\left[\mathbf{E}\left[f \mid \mathscr{F}_{A^{c}}\right]^{2} - \mathbf{E}\left[f \mid \mathscr{F}_{(A \cup B)^{c}}\right]^{2}\right]$$
$$= \mathbf{E}\left[\left(\mathbf{E}\left[f \mid \mathscr{F}_{A^{c}}\right] - \mathbf{E}\left[f \mid \mathscr{F}_{(A \cup B)^{c}}\right]\right)^{2}\right].$$

### So, what are we going to do?

With quite a lot of work for both items,

(1') 
$$\mathbf{P}\Big[|\mathscr{S}_n \cap B_r/3| > c \,\mathbf{E}|\mathscr{S}_r| \mid \mathscr{S}_n \cap B_r \neq \emptyset = \mathscr{S}_n \cap W\Big] \ge c > 0.$$

(2)  $\mathbf{P}[|\mathscr{S}_n(r)| = k] \leq g(k) \mathbf{P}[|\mathscr{S}_n(r)| = 1]$ , with sub-exponential g(k).

We could repeat (1') for many *r*-boxes only if "not enough points in one box" meant "we found nothing in that box".

So, take an independent random dilute sample:  $\mathbf{P}[x \in \mathcal{R}] = 1/\mathbf{E}|\mathscr{S}_r|$  i.i.d. Then,  $|\mathscr{S}_n \cap B_r/3|$  is small  $\implies \mathcal{R} \cap \mathscr{S}_n \cap B_r/3 = \emptyset$  is likely, and  $|\mathscr{S}_n \cap B_r/3|$  is large  $\implies \mathcal{R} \cap \mathscr{S}_n \cap B_r/3 \neq \emptyset$  is likely.

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But  $\mathbf{P}\left[\mathscr{S}_n \neq \emptyset = \mathcal{R} \cap \mathscr{S}_n \mid |\mathscr{S}_n(r)| = k\right]$  is still problematic conditioning.

A strange large deviations lemma solves the issue.

#### The strange large deviation lemma

Suppose  $X_i, Y_i \in \{0, 1\}$ , i = 1, ..., n, and that  $\forall J \subset [n]$  and  $\forall i \in [n] \setminus J$ 

$$\mathbf{P}\left[Y_i=1 \mid \forall_{j\in J}Y_j=0\right] \geqslant c \mathbf{P}\left[X_i=1 \mid \forall_{j\in J}Y_j=0\right].$$

Then

$$\mathbf{P}\Big[\forall_i Y_i = \mathbf{0}\Big] \leqslant c^{-1} \mathbf{E}\Big[\exp\Big(-(c/e)\sum_i X_i\Big)\Big].$$

We use this with  $X_j := 1_{\{\mathscr{S} \cap B_j \neq \emptyset\}}$  and  $Y_j := 1_{\{\mathscr{S} \cap B_j \cap \mathcal{R} \neq \emptyset\}}$ .

**Proof:** Instead of sequential scan, average everything together. Choose  $J \subset [n]$  randomly, Bernoulli(1-p). Get  $\mathbf{E}[Y p^Y] \ge c \mathbf{E}[X p^{Y+1}]$ .

So,  $\mathbf{E}[Z] \ge 0$ , where  $Z := (Y - c p X) p^Y$ . Choose  $p := e^{-1}$ . Maximize Z over Y, and get the bound  $Z \le \exp(-1 - c X/e)$ . Altogether,  $c e^{-1} \mathbf{P}[Y = 0 < X] \le \mathbf{E}[1_{X>0} \exp(-1 - c X/e)]$ , and done.

### Final result for the spectral sample

If  $r \in [1, n]$ , then  $\{|\mathscr{S}_n| < \mathbf{E}|\mathscr{S}_r|\}$  is basically equivalent to being contained inside some  $r \times r$  sub-square:

$$\mathbf{P}\left[0 < |\mathscr{S}_n| < \mathbf{E}|\mathscr{S}_r|\right] \asymp \alpha_4(r,n)^2 \left(\frac{n}{r}\right)^2.$$

In particular, on the triangular lattice  $\Delta$ ,

 $\mathbf{P}\big[0 < |\mathscr{S}_n| < \lambda \, \mathbf{E}|\mathscr{S}_n| \big] \asymp \lambda^{2/3}.$ 

The *scaling limit* of  $\mathscr{S}_n$  is a conformally invariant Cantor-set with Hausdorffdimension 3/4.

GPS (2010-12) proves that the scaling limit of dynamical percolation exists as a Markov process; for mesh 1/n the time-scale is  $tn^{-3/4+o(1)}$ . The above implies that this process is ergodic, with correlations decaying as  $t^{-2/3}$ .

Question 1: Can one build similar proofs for other Boolean functions?

Question 2: Self-similarity of  $Piv_n$  and  $\mathscr{S}_n$  is a lot of restriction on these random sets. And it's not only because of conformal invariance: Gil Kalai noticed that the spectral sample of recursive 3-wise majority is the leaves of a GW-tree! This phenomenon might be somehow general:

Influence-Entropy conjecture [Friedgut-Kalai 1996]: For some universal constant C, for any Boolean function f,

$$\begin{split} \mathbf{SpecEnt}(f) &:= \sum_{S \subset [n]} \widehat{f}(S)^2 \log \frac{1}{\widehat{f}(S)^2} \leqslant C \times \\ &\times \mathbf{Influence}(f) := \mathbf{E}|\mathscr{S}_f| = \mathbf{E}|\mathrm{Piv}_f| = \sum_{S \subset [n]} \widehat{f}(S)^2 |S| \,. \end{split}$$

I.e., there is no  $\log$  factor in the entropy as it would be in uniform.

I think I can do it for  $Piv_n$ , but not enough independence is known in  $\mathscr{S}_n$ .