

Upper tail decay of KPZ models with Brownian initial conditions

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Abstract

In this paper we consider the limiting distribution of KPZ growth models with random but not stationary initial conditions introduced in [6]. The one-point distribution of the limit is given in terms of a variational problem. By directly studying it, we deduce the right tail asymptotic of the distribution function. This gives a rigorous proof and extends the results obtained by Meerson and Schmidt in [18].

1 Introduction

The Kardar–Parisi–Zhang (KPZ) universality class of stochastic growth models in one dimension are described by a stochastically growing interface parameterized by a height function. For general initial conditions, the one-point distribution of the large time limit can be written in terms of a variational problem. The ingredients are the (scaled) initial condition and the \mathcal{A}_2 process, \mathcal{A}_2 . The latter arises as the limiting interface process when the macroscopic geometry of the interface in the law of large numbers is curved. It was discovered in the work of Prähofer and Spohn [23] and described by its finite-dimensional distribution. Soon after, Johansson showed weak convergence of the discrete polynuclear growth model to the \mathcal{A}_2 process [16]. In the same paper, a first variational formula appeared (see Corollary 1.3 of [16])

$$F_{\text{GOE}}(2^{2/3}s) = \mathbf{P} \left(\sup_{t \in \mathbb{R}} (\mathcal{A}_2(t) - t^2) \leq s \right) \quad (1.1)$$

where \mathcal{A}_2 is the \mathcal{A}_2 process and F_{GOE} is the GOE Tracy–Widom distribution function discovered in random matrix theory [27]. Formula (1.1) corresponds to the flat initial condition as F_{GOE} is the limiting distribution of the corresponding rescaled interface.

Later, variational formulas describing the one-point distributions for some special initial conditions appeared in several papers, see for instance [3, 24–26]. The first study of a large class of initial conditions, including random initial conditions, is the paper of Corwin, Liu and Wang [8]. In a last passage percolation model they showed the

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convergence of the one-point distribution to a probability distribution expressed by the variational formula

$$\mathbf{P} \left(\sup_{t \in \mathbb{R}} \{h_0(t) + \mathcal{A}_2(t) - t^2\} \leq s \right) \quad (1.2)$$

where h_0 is the scaling limit of the initial height profile. Shortly after, Remenik and Quastel in [26] asked and answered the question how much discrepancy from the perfectly flat initial condition would be allowed to still see the GOE Tracy–Widom distribution for the KPZ equation. In their paper the variational representation plays an important role. The variational formula approach is proved to be useful since it allows one to go beyond the use of exact formulas and to show, for instance, universal limiting distribution for a flat but tilted profile [11].

Building on [8], Chhita, Ferrari and Spohn derived a variational formula which describes the limiting distribution for random initial conditions which scale to a Brownian motion with the result [6]

$$F^{(\sigma)}(s) = \mathbf{P} \left(\sup_{t \in \mathbb{R}} \left\{ \sqrt{2}\sigma B(t) + \mathcal{A}_2(t) - t^2 \right\} \leq s \right) \quad (1.3)$$

where B is a standard two-sided Brownian motion independent of the Airy_2 process \mathcal{A}_2 . This distribution has two special cases which could be analyzed using exact formulas, namely $\sigma = 0$ is the flat case and it reduces to (1.1) whereas $\sigma = 1$ corresponds to the stationary initial condition for the model, so that $F^{(1)}(s)$ is the Baik–Rains distribution [4].

The characterization through a variational formula is tightly related to the question of universality. In the framework of this paper, the key universal ingredient is the Airy_2 process which is a projection of a more general space-time random process. The study of this process started with the discovery of the KPZ fixed point by Matetski, Quastel and Remenik [17], for further properties see [5, 20, 21], and continued with the description of the full space-time process called the Airy sheet or also directed landscape by Dauvergne, Ortman and Virág [10], see also [19].

Deducing concrete information from a variational formula is however not always an easy task. For example, given (1.3), it is not clear what are the tails of the distribution. They have only been known for a long time in the cases $\sigma = 0$ and $\sigma = 1$, because these distributions had other representations, see e.g. [22]. Meerson and Schmidt considered the $F^{(\sigma)}$ distribution in [18] and they deduced the correct right tail behavior by a physically motivated but non-rigorous method. They found that $\ln(1 - F^{(\sigma)}(s)) \sim -\frac{4}{3} \frac{1}{\sqrt{1+3\sigma^4}} s^{3/2}$ for $s \gg 1$. They also performed large scale simulations on the exclusion process confirming their finding.

In this paper we give a rigorous proof of the asymptotics and extend the results of [18] by obtaining upper and lower bound on the prefactor in front of the stretched exponential decay, see Theorem 1.1. The upper tail distribution is governed by the maximal value of $\sqrt{2}\sigma B(t) - t^2$. This fact holds already for non-random initial conditions. For instance, in the case of (1.1), the tail behaviour of $1 - F_{\text{GOE}}(2^{2/3}s)$ matches that of $1 - F_{\text{GUE}}(s)$ up to the exponential scale as the maximal value of $-t^2$ is obtained at $t = 0$. The same was shown for another simple function h_0 in (1.2) as noticed in [28]. To make this point explicit, Theorem 1.3 gives the tail decay for a generic non-random initial condition. One important ingredient for the proof of Theorems 1.1 and 1.3 is the observation that, for

all $\delta > 0$, the tail distribution of

$$\mathbf{P} \left(\sup_{t \in \mathbb{R}} (\mathcal{A}_2(t) - \delta t^2) > s \right) \quad (1.4)$$

is, in the exponential scale, independent of δ , see Theorem 1.2 for a detailed statement.

For a given realization of the Brownian motion, if the supremum in (1.1) is taken over a finite interval instead of \mathbb{R} , the distribution we are considering also has a Fredholm determinant expression with a kernel depending on the Brownian motion [9]. This representation is however not directly applicable when taking the limit as the finite interval approaches \mathbb{R} . For this purpose, it is better to use the kernel given in terms of hitting times noticed first in [26]. That representation works well provided that the function h_0 is lower than a parabola with a prefactor $3/4$, while in our application we need to get close to 1. The explicit kernel representation has some intrinsic technical issues as confirmed also in the simplest case of a hat-shaped h_0 [28]. Our method is mainly probabilistic and avoids the computation of a correlation kernel and its asymptotic analysis.

The second issue that we had to deal with was that the density of the maximum of a Brownian motion with parabolic drift studied first by Groeneboom [13], see also [14] for explicit formulas, contains a term with a linear combination of Airy Ai and Airy Bi functions. The leading term is however coming from a subtle cancellation and it does not follow from the naive asymptotic of the Airy functions. Fortunately, we could avoid this issue by using an integral representation discovered by Janson, Louchard and Martin-Löf in [15], which we carefully analyzed asymptotically, see Proposition 3.1.

The paper is organized as follows. We state the main results in the rest of the introduction. We first prove Theorem 1.2 on the upper tail of the supremum of the Airy₂ process minus a parabola with arbitrary coefficient in Section 2. Then Section 3 is about the asymptotic of the supremum of the Brownian motion minus a parabola. Section 4 proves Theorem 1.1 on the right tail of the limiting distribution $F^{(\sigma)}$ for Brownian initial conditions. The proof of Theorem 1.3 about the case of general deterministic initial conditions is given in Section 5.

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Main results

Theorem 1.1. *Let $\sigma > 0$ be fixed. For s large enough, the right tail of $F^{(\sigma)}(s)$ satisfies*

$$C_1 s^{-3/4} e^{-\frac{4}{3} \frac{1}{\sqrt{1+3\sigma^4}} s^{3/2}} \leq 1 - F^{(\sigma)}(s) \leq C_2 s^{3/4} \ln(s) e^{-\frac{4}{3} \frac{1}{\sqrt{1+3\sigma^4}} s^{3/2}} \quad (1.5)$$

for some constants C_1, C_2 independent of s .

The $s^{-3/4}$ behavior of the prefactor of the lower bound seems to be the correct one. Indeed, for $\sigma = 0$ the prefactor is $s^{-3/4}/(4\sqrt{2\pi})$ (see (1.8) below) and for $\sigma = 1$ it is given by¹ $s^{-3/4}/\sqrt{\pi}$. In Proposition 4.1 and 4.2 we give expressions of the σ -dependence of C_1, C_2 .

As it could be expected by the definition of the $F^{(\sigma)}$ distribution (1.3), the upper tail behaviour of the $F^{(\sigma)}$ distribution is related to tail of the maximum of the Airy_2 process minus a parabola. The variational formula (1.1) was proved in [16]. In the next result we generalize [16] in the sense that we compute the tail decay for the supremum with parabola which can have any coefficient between 0 and 1.

Theorem 1.2. *Let $(\mathcal{A}_2(t))_{t \in \mathbb{R}}$ denote the Airy_2 process. There is a constant $C > 0$ such that for all $c \in (0, 1)$*

$$1 - F_{\text{GUE}}(s) \leq \mathbf{P} \left(\sup_{t \in \mathbb{R}} (\mathcal{A}_2(t) - (1-c)t^2) > s \right) \leq C \frac{\ln(s/(1-c))}{s^{3/4}\sqrt{1-c}} e^{-\frac{4}{3}s^{3/2}} \quad (1.6)$$

holds as $s \rightarrow \infty$ where the constant C is independent of s .

This result might be compared to what has been proven for the KPZ equation with narrow wedge initial condition at finite but large time. In that case, it is known that the decay is exponential with $s^{3/2}$ power, but without a more precise information on the coefficient, see Proposition 4.2 of [7].

The lower bound in (1.6) is obvious by taking $t = 0$ instead of the supremum. Its asymptotic expansion follows from (1) and (25) of [2], namely

$$1 - F_{\text{GUE}}(s) = \frac{1}{16\pi s^{3/2}} e^{-\frac{4}{3}s^{3/2}} (1 + \mathcal{O}(s^{-3/2})) \quad (1.7)$$

as $s \rightarrow \infty$. We remark that for $c \leq 0$ the upper bound in (1.6) is trivial as well, since

$$\mathbf{P} \left(\sup_{t \in \mathbb{R}} (\mathcal{A}_2(t) - (1-c)t^2) > s \right) \leq \mathbf{P} \left(\sup_{t \in \mathbb{R}} (\mathcal{A}_2(t) - t^2) > s \right) \sim \frac{1}{4\sqrt{2\pi}s^{3/4}} e^{-\frac{4}{3}s^{3/2}} \quad (1.8)$$

where we used (1.1) and the $x \rightarrow \infty$ tail asymptotic

$$1 - F_{\text{GOE}}(x) \sim \frac{e^{-\frac{2}{3}x^{3/2}}}{4\sqrt{\pi}x^{3/4}}. \quad (1.9)$$

Theorem 1.3. *Let $h_0 : \mathbb{R} \rightarrow \mathbb{R}$ be a function satisfying $h_0(t) \leq A + (1-\varepsilon)t^2$ for all $t \in \mathbb{R}$, for some constants $A \in \mathbb{R}$ and $\varepsilon > 0$. Let $\kappa(h_0) = \sup_{t \in \mathbb{R}} \{h_0(t) - t^2\}$ and let $M > 0$ be large enough so that $h_0(t) \leq \kappa(h_0) + (1 - \frac{\varepsilon}{2})t^2$ for all $|t| \geq M$. Then there are positive real constants C_1 and C_2 which do not depend on the function h_0 and s , such that for s large enough*

$$C_1 \frac{e^{-\frac{4}{3}(s-\kappa(h_0))^{3/2}}}{(s-\kappa(h_0))^{3/2}} \leq \mathbf{P} \left(\sup_{t \in \mathbb{R}} \{h_0(t) + \mathcal{A}_2(t) - t^2\} \geq s \right) \leq C_2 M \frac{e^{-\frac{4}{3}(s-\kappa(h_0))^{3/2}}}{(s-\kappa(h_0))^{1/4}}. \quad (1.10)$$

¹For $\sigma = 0$, the prefactor is obtained from equations (1) and (25), (26) of [2], except that in (26) there is a typo, namely $x^{-3/2}$ should be $x^{-3/4}$. It can be also easily obtained from the Fredholm determinant representation of F_{GOE} in [12]. For $\sigma = 1$ the distribution is the Baik–Rains distribution, given in Definition 2 of [4]. The prefactor easily follows using (2.3), (2.6) of [4], as well as (26) of [2].

2 Supremum of the Airy_2 process minus a parabola

The aim of this section is to prove Theorem 1.2 about the upper tail behaviour of the Airy_2 process minus a parabola with arbitrary coefficient. The first ingredient is a simple bound on the supremum of the Airy_2 process over a finite interval.

Lemma 2.1. *There is a explicit constant C such that for all $a > 0$*

$$\mathbf{P}\left(\sup_{t \in [0, a]} \mathcal{A}_2(t) > s\right) \leq C \frac{e^{-\frac{4}{3}(s-a^2)^{3/2}}}{(s-a^2)^{3/4}} \quad (2.1)$$

holds if s is large enough. From this we get the bound

$$\mathbf{P}\left(\sup_{t \in [0, a]} \mathcal{A}_2(t) > s\right) \leq C' \frac{a}{s^{1/4}} e^{-\frac{4}{3}s^{3/2}} \quad (2.2)$$

for large s with some other constant C' .

Proof. The probability on the left-hand side of (2.1) can be upper bounded as

$$\begin{aligned} \mathbf{P}\left(\sup_{t \in [0, a]} \mathcal{A}_2(t) > s\right) &\leq \mathbf{P}\left(\sup_{t \in [0, a]} (\mathcal{A}_2(t) - t^2) > s - a^2\right) \\ &\leq \mathbf{P}\left(\sup_{t \in \mathbb{R}} (\mathcal{A}_2(t) - t^2) > s - a^2\right) \\ &= 1 - F_{\text{GOE}}(2^{2/3}(s - a^2)) \end{aligned} \quad (2.3)$$

where we extended the range of t in the second inequality and used (1.1) in the last equality above which is the result of [16]. The proof of the first inequality (2.1) with $C = 1/(4\sqrt{2\pi})$ is completed by applying the upper tail behaviour of the F_{GOE} distribution (1.9).

For (2.2), we divide the interval $[0, a]$ into pieces of length $1/\sqrt{s}$ and by the union bound we get

$$\begin{aligned} \mathbf{P}\left(\sup_{t \in [0, a]} \mathcal{A}_2(t) > s\right) &\leq \sum_{k=1}^{a\sqrt{s}} \mathbf{P}\left(\sup_{t \in [(k-1)/\sqrt{s}, k/\sqrt{s}]} \mathcal{A}_2(t) > s\right) \\ &\leq C a\sqrt{s} \frac{e^{-\frac{4}{3}(s-1/s)^{3/2}}}{(s-1/s)^{3/4}} \\ &\leq C' \frac{a}{s^{1/4}} e^{-\frac{4}{3}s^{3/2}} \end{aligned} \quad (2.4)$$

for some constant C' where we applied stationarity of the Airy_2 process and (2.1) in the second inequality and the bound $-(s-1/s)^{3/2} \leq -s^{3/2} + 2$ for all $s \geq 1$ in the third inequality above. \square

With this lemma we are ready to prove Theorem 1.2.

Proof of Theorem 1.2. To bound the probability that the Airy_2 process remains below a parabola, consider an increasing sequence $x_0 = 0 < x_1 < x_2 < \dots$ to be a partition of

\mathbb{R}_+ to be specified later. Then we have by symmetry and the union bound

$$\begin{aligned}
\mathbf{P} \left(\sup_{t \in \mathbb{R}} (\mathcal{A}_2(t) - (1-c)t^2) > s \right) &\leq 2\mathbf{P} \left(\cup_{k=0}^{\infty} \{ \exists t \in [x_k, x_{k+1}] : \mathcal{A}_2(t) > s + (1-c)t^2 \} \right) \\
&\leq 2 \sum_{k=0}^{\infty} \mathbf{P} \left(\exists t \in [x_k, x_{k+1}] : \mathcal{A}_2(t) > s + (1-c)t^2 \right) \\
&\leq 2 \sum_{k=0}^{\infty} \mathbf{P} \left(\exists t \in [x_k, x_{k+1}] : \mathcal{A}_2(t) > s + (1-c)x_k^2 \right) \\
&\leq 2C \sum_{k=0}^{\infty} \frac{e^{-\frac{4}{3}(s+(1-c)x_k^2 - (x_{k+1}-x_k)^2)^{3/2}}}{(s + (1-c)x_k^2 - (x_{k+1} - x_k)^2)^{3/4}}
\end{aligned} \tag{2.5}$$

where we decreased the barrier which $\mathcal{A}_2(t)$ has to reach in $[x_k, x_{k+1}]$ in the third inequality, while in the fourth inequality we used Lemma 2.1 and the translation invariance of the Airy_2 process.

The $k = 0$ term in the sum on the right-hand side of (2.5) is $e^{-\frac{4}{3}(s-x_1^2)^{3/2}}/(s-x_1^2)^{3/4}$, hence if we choose $x_1 = 1/\sqrt{s}$, then it is still bounded by a constant multiplied by $e^{-\frac{4}{3}s^{3/2}}/s^{3/4}$. The rest of the sequence is chosen so that it satisfies

$$(x_{k+1} - x_k)^2 = \frac{1}{4}(1-c)x_k^2 \tag{2.6}$$

for $k = 1, 2, \dots$ which is the geometric choice $x_{k+1} = \gamma x_k$ with $\gamma = 1 + \frac{\sqrt{1-c}}{2}$. With this sequence, the right-hand side of (2.5) can be bounded as

$$\begin{aligned}
\mathbf{P} \left(\sup_{t \in \mathbb{R}} (\mathcal{A}_2(t) - (1-c)t^2) > s \right) &\leq 2C \frac{e^{-\frac{4}{3}(s-\frac{1}{s})^{3/2}}}{(s-\frac{1}{s})^{3/4}} + 2C \sum_{k=1}^{\infty} \frac{e^{-\frac{4}{3}(s+\frac{3}{4}(1-c)x_k^2)^{3/2}}}{s^{3/4}} \\
&\leq 2C' \frac{e^{-\frac{4}{3}s^{3/2}}}{s^{3/4}} + 2C' \frac{e^{-\frac{4}{3}s^{3/2}}}{s^{3/4}} \sum_{k=1}^{\infty} e^{-\frac{4}{3}(\frac{3}{4}(1-c)\gamma^{2(k-1)}s^{-1})^{3/2}}
\end{aligned} \tag{2.7}$$

where we used the inequality $(a+b)^{3/2} \geq a^{3/2} + b^{3/2}$ in the last step. The sum on the right-hand side of (2.7) can be upper bounded using Lemma 2.2 below with $\alpha = \gamma^3$ and $\beta = \sqrt{3}(1-c)^{3/2}/(2s^{3/2})$ as

$$\sum_{k=1}^{\infty} e^{-\frac{\sqrt{3}(1-c)^{3/2}}{2s^{3/2}}\gamma^{3(k-1)}} \leq \frac{\ln \left(1 + \frac{2\gamma^3 s^{3/2}}{\sqrt{3}(1-c)^{3/2}} \right)}{3 \ln \gamma} \leq \tilde{C} \frac{\ln(s/(1-c))}{\sqrt{1-c}} \tag{2.8}$$

for s large enough with some \tilde{C} which does not depend on c . In the last inequality above we used that $\ln \gamma \sim \frac{1}{2}\sqrt{1-c}$ as $c \rightarrow 1$. The inequalities (2.7) and (2.8) together prove (1.6). \square

Lemma 2.2. *Let $\alpha > 1$ and $\beta > 0$. Then*

$$\sum_{k=0}^{\infty} e^{-\beta\alpha^k} \leq \frac{\ln(1 + \alpha/\beta)}{\ln \alpha} \tag{2.9}$$

Proof. We start by bounding the sum by an integral as

$$\sum_{k=0}^{\infty} e^{-\beta\alpha^k} \leq \int_{-1}^{\infty} dx e^{-\beta\alpha^x}. \quad (2.10)$$

The change of variables $w = \beta\alpha^x$ gives $\frac{dw}{dx} = w \ln \alpha$ so that the right-hand side of (2.10)

$$\int_{-1}^{\infty} dx e^{-\beta\alpha^x} = \frac{1}{\ln \alpha} \int_{\beta/\alpha}^{\infty} dw \frac{e^{-w}}{w} =: \frac{1}{\ln \alpha} E_1(\beta/\alpha). \quad (2.11)$$

Equation (5.1.20) of [1] provides a bound on the exponential integral function E_1 , namely $E_1(x) \leq e^{-x} \ln(1 + 1/x)$ for $x > 0$. Thus

$$\sum_{k=0}^{\infty} e^{-\beta\alpha^k} \leq \frac{1}{\ln \alpha} e^{-\beta/\alpha} \ln(1 + \alpha/\beta) \quad (2.12)$$

which gives the claimed bound since $\alpha, \beta > 0$. □

3 Supremum of Brownian motion minus a parabola

Proposition 3.1. *Let $G(x) = \mathbf{P}(\max_{t \in \mathbb{R}}(B(t) - \frac{1}{2}t^2) \geq x)$ where $B(t)$ is a standard two-sided Brownian motion. Then, as $x \rightarrow \infty$ we have*

$$G(x) = 3^{-1/2} e^{-\frac{4}{3}\sqrt{\frac{2}{3}}x^{3/2}} (1 + \mathcal{O}(x^{-1/4})) \quad (3.1)$$

as well as

$$-\frac{d}{dx} G(x) = \frac{2\sqrt{2}}{3} e^{-\frac{4}{3}\sqrt{\frac{2}{3}}x^{3/2}} \sqrt{x} (1 + \mathcal{O}(x^{-1/4})). \quad (3.2)$$

Consequently, for any $c > 0$ the density of the random variable $\max_{t \in \mathbb{R}}(B(t) - ct^2)$ satisfies

$$f_c(x) := -\frac{d}{dx} G((2c)^{1/3}x) = \frac{4}{3} \sqrt{c} e^{-\frac{4}{3}\sqrt{\frac{4}{3}}c^{1/3}x^{3/2}} \sqrt{x} (1 + \mathcal{O}(x^{-1/4})) \quad (3.3)$$

as $x \rightarrow \infty$.

The proof is given below. The distribution function $G(x)$ is written as a contour integral in Lemma 3.5 of [15] as follows:

$$G(x) = \frac{1}{2i} \int_{\gamma} dz \frac{\text{Hi}(z)}{\text{Ai}(z)} \text{Ai}(z + 2^{1/3}x) \quad (3.4)$$

where γ is a path passing to the right of all zeroes of the Airy function Ai from $-\infty$ to ∞ . The function Hi is defined by (see (10.4.44) of [1])

$$\text{Hi}(z) = \pi^{-1} \int_0^{\infty} dt e^{-t^3/3 + zt}. \quad (3.5)$$

In [15], the contour γ in (3.4) was chosen to come from $e^{-i\theta}\infty$ and arrive to $e^{i\theta}\infty$ with θ slightly larger than $\pi/2$. The reason why this contour can be deformed to the vertical one is the following. Lemma A.1 of [15] was used to argue for convergence of the integral (3.4): they showed the decay of the ratio of the two Airy Ai functions, together with a

bound on Hi for contours with angle more than $\pi/2$. The bound in Lemma 3.3 is enough to get the convergence also for vertical contours.

For the proof of (3.1), the asymptotic of the Airy and Hi functions will be needed. By (10.4.90) of [1], for x real,

$$\text{Hi}(x) \sim \pi^{-1/2} x^{-1/4} e^{\frac{2}{3}x^{3/2}} \text{ as } x \rightarrow \infty. \quad (3.6)$$

For our purposes, we need also the asymptotic behavior of $\text{Hi}(z)$ for complex-valued z close to $2^{1/3}x/3$ which we state below and prove later in this section.

Lemma 3.2. *Let z be such that $|\arg(z)| < \pi/3$. Then for large z we have the asymptotic behavior*

$$\text{Hi}(z) e^{-\frac{2}{3}z^{3/2}} = \pi^{-1/2} z^{-1/4} + \mathcal{O}(|z|^{-1/2}). \quad (3.7)$$

Lemma 3.3. *Let $\theta \in [\pi/2, 3\pi/2]$ and $x \in \mathbb{R}$. Then for all $y \geq 0$,*

$$|\text{Hi}(x + e^{i\theta}y)| \leq \text{Hi}(x). \quad (3.8)$$

Proof. By the definition (3.5),

$$\text{Hi}(x + e^{i\theta}y) = \pi^{-1} \int_0^\infty dt e^{-t^3/3} e^{xt} e^{e^{i\theta}yt} \quad (3.9)$$

where $|e^{e^{i\theta}yt}| \leq 1$ for all $\theta \in [\pi/2, 3\pi/2]$ and $t \geq 0$. Since $e^{-t^3/3} e^{xt}$ is positive, the absolute value of the integral in (3.9) can be upper bounded by the integral of $e^{-t^3/3} e^{xt}$ which yields (3.8). \square

Proof of Proposition 3.1. Let us first prove (3.1). In order to estimate $G(x)$ for large x , we use the integral representation (3.4). The integration contour is chosen to be vertical $\gamma = \frac{2^{1/3}}{3}x + i\mathbb{R}$. If $x > 0$, then for $z \in \gamma$, we have $\arg(z) \in [-\pi/2, \pi/2]$. Hence we can use the asymptotics of the Airy function

$$\text{Ai}(z) = \frac{1}{2}\pi^{-1/2} z^{-1/4} e^{-\frac{2}{3}z^{3/2}} (1 + O(1/z)) \text{ for } |\arg(z)| < \pi, \quad (3.10)$$

see (10.4.59) of [1]. Let us parameterize the path γ as $z = \frac{2^{1/3}}{3}x + i2^{1/3}xv$, $v \in \mathbb{R}$.

Contribution for $|v| > 1/3$. Using Lemma 3.3 and 3.2 for the Hi function and the asymptotic (3.10) on the Airy functions, the contribution for $|v| > 1/3$ is bounded by

$$Cx^{3/4} \int_{1/3}^\infty dv e^{-x^{3/2}g(v)}, \quad g(v) = \frac{2\sqrt{2}}{9\sqrt{3}} \text{Re}[(4 + 3iv)^{3/2} - (1 + 3iv)^{3/2} - 1] \quad (3.11)$$

for some constant C . Notice that for $a \geq 0$,

$$\frac{d \text{Re}((a + iv)^{3/2})}{da} = -\frac{3}{2} \text{Im} \sqrt{a + iv} \quad (3.12)$$

which is an increasing function of a . This implies that $g(v)$ is monotone increasing in $v \geq 0$ with $g(v) \sim \sqrt{v}$ as $v \rightarrow \infty$. We get that the leading behavior of the integral is bounded by $C'e^{-x^{3/2}g(1/3)}$ with $g(1/3) - g(0) < 0$, for some other constant C' . Thus this contribution is vanishing with respect to the leading one computed below.

Contribution for $|v| \leq 1/3$. For the leading contribution, we use the asymptotic expansion of Lemma 3.2 and (3.10), with the result that the contribution of $G(x)$ coming from the integral over $|v| \leq 1/3$ is given by

$$\begin{aligned} 2^{-2/3}x \int_{|v| \leq 1/3} dv \frac{\text{Hi}\left(\frac{2^{1/3}}{3}x + i2^{1/3}xv\right)}{\text{Ai}\left(\frac{2^{1/3}}{3}x + i2^{1/3}xv\right)} \text{Ai}\left(\frac{4 \cdot 2^{1/3}}{3}x + i2^{1/3}xv\right) \\ = \frac{x^{3/4}}{\sqrt{\pi}2^{3/4}} \int_{|v| \leq 1/3} dv \frac{e^{x^{3/2}h(v)}}{(4/3 + iv)^{1/4}} (1 + \mathcal{O}(1/x^{1/4})) \end{aligned} \quad (3.13)$$

with $h(v) = \frac{4}{3}\sqrt{2}((1/3 + iv)^{3/2} - (4/3 + iv)^{3/2}/2)$. Further, one notices that $\text{Re}(h(v))$ is strictly increasing for $v < 0$ and strictly decreasing for $v > 0$ with a quadratic approximation for small v given by

$$h(v) = -\frac{4\sqrt{2}}{3\sqrt{3}} - \frac{3\sqrt{3}}{4\sqrt{2}}v^2 + \mathcal{O}(v^3). \quad (3.14)$$

Then, using standard steep descent analysis, we get (3.1).

The proof of (3.2) is similar. The only difference is that we need to replace the asymptotic expansion of $\text{Ai}(z + 2^{1/3}x)$ with the one of $2^{1/3} \text{Ai}'(z + 2^{1/3}x)$. By (10.4.61) of [1], we have $\text{Ai}'(z) = -\frac{1}{2}\pi^{-1/2}z^{1/4}e^{-2z^{3/2}/3}(1 + \mathcal{O}(1/z))$. Thus in the asymptotic analysis we need to replace $z^{-1/4}$ with $-2^{1/3}z^{1/4}$. This gives the claimed result.

By replacing t by $(2c)^{-2/3}t$ and using Brownian rescaling,

$$\mathbf{P}\left(\max_{t \in \mathbb{R}} (B(t) - ct^2) \geq x\right) = \mathbf{P}\left(\max_{t \in \mathbb{R}} (B(t) - \frac{1}{2}t^2) \geq x(2c)^{1/3}\right) = G((2c)^{1/3}x), \quad (3.15)$$

hence (3.3) follows from (3.2) by substitution. \square

Now we prove the claimed asymptotic expansion of Hi .

Proof of Lemma 3.2. By symmetry with respect to the real axis ($z \mapsto \bar{z}$), consider z with $\arg(z) \in [0, \pi/3)$. We parameterize $z = re^{i\theta}$ with $r > 0$ and $\theta \in [0, \pi/3)$. We introduce the change of variables $t = \sqrt{z} + u$ which yields

$$-\frac{1}{3}t^3 + zt = \frac{2}{3}z^{3/2} - \frac{1}{3}u^3 - \sqrt{z}u^2 \quad (3.16)$$

and we get

$$\text{Hi}(z) e^{-2z^{3/2}/3} = \pi^{-1} \int_{-\sqrt{z}}^{\infty} du e^{-u^3/3 - \sqrt{z}u^2}. \quad (3.17)$$

For large values of $|z|$, the leading contribution comes from a neighborhood of 0. Consider the integration contour $\Gamma = \Gamma_1 \vee \Gamma_2$ where $\Gamma_1 = \{-\sqrt{z} + iy, 0 \leq y \leq \text{Im}(\sqrt{z})\}$ and $\Gamma_2 = \{x, -\text{Re}(\sqrt{z}) \leq x < \infty\}$.

Consider first the contribution on the contour Γ_1 . Let $f(u) = \text{Re}(-u^3/3 - \sqrt{z}u^2)$. Then we have

$$f(-\sqrt{z}) = -\frac{2}{3} \text{Re}(z^{3/2}) = -\frac{2}{3}r^{3/2} \cos(3\theta/2) < 0 = f(0). \quad (3.18)$$

Setting $u = -\sqrt{z} + iy = -\sqrt{r} \cos(\theta/2) - i\sqrt{r} \sin(\theta/2) + iy$, we obtain

$$\operatorname{Re}(-u^3/3 - \sqrt{z}u^2) = \operatorname{const} - \sqrt{r} \sin(\theta)y, \quad (3.19)$$

which is decreasing in y . Thus the contribution of the integral of (3.17) over Γ_1 can be bounded by the maximum of the integrand times the length of the contour, that is, by $\pi^{-1} \operatorname{Im}(\sqrt{z})e^{f(-\sqrt{z})} \leq Ce^{-\frac{2}{3}r^{3/2} \cos(3\theta/2)}$.

Next we focus on the contribution over Γ_2 . We have, for all $u \geq -\sqrt{r} \cos(\theta/2) = -\operatorname{Re}(\sqrt{z})$, the bound

$$\operatorname{Re}(-u^3/3 - \sqrt{z}u^2) = -\frac{1}{3}u^3 - \sqrt{r} \cos(\theta/2)u^2 \leq -\frac{2}{3}\sqrt{r} \cos(\theta/2)u^2. \quad (3.20)$$

For any $\delta > 0$, which can be chosen later as a function of r , the tails of the Gaussian integral gives

$$\left| \pi^{-1} \int_{\Gamma_2 \setminus \{|u| \leq \delta\}} du e^{-u^3/3 - \sqrt{z}u^2} \right| \leq Ce^{-\frac{2}{3}\sqrt{r} \cos(\theta/2)\delta^2}. \quad (3.21)$$

Next, the local contribution is close to the integral with only the quadratic term. Indeed, using $|e^x - 1| \leq |x|e^{|x|}$, we have

$$\begin{aligned} \left| \pi^{-1} \int_{\{|u| \leq \delta\}} du e^{-u^3/3 - \sqrt{z}u^2} - \pi^{-1} \int_{\{|u| \leq \delta\}} du e^{-\sqrt{z}u^2} \right| \\ \leq \pi^{-1} \int_{\{|u| \leq \delta\}} du \left| e^{-u^3/3 - \sqrt{z}u^2} \right| \frac{|u|^3}{3} = \mathcal{O}(\delta^3). \end{aligned} \quad (3.22)$$

Extending the integration in the Gaussian integral to \mathbb{R} , we only make a small error, namely

$$\left| \pi^{-1} \int_{\{|u| \leq \delta\}} du e^{-\sqrt{z}u^2} - \pi^{-1} \int_{\mathbb{R}} du e^{-\sqrt{z}u^2} \right| \leq \mathcal{O}(e^{-\delta^2 \sqrt{r} \cos(\theta/2)}). \quad (3.23)$$

Finally, the Gaussian integral can be computed explicitly as

$$\pi^{-1} \int_{\mathbb{R}} du e^{-\sqrt{z}u^2} = \pi^{-1/2} z^{-1/4}. \quad (3.24)$$

Combining all these bounds, we get

$$\operatorname{Hi}(z) e^{-\frac{2}{3}z^{3/2}} = \pi^{-1/2} z^{-1/4} + \mathcal{O}(\delta^3, \sqrt{r}e^{-\frac{2}{3}r^{3/2} \cos(3\theta/2)}, e^{-\frac{2}{3}\sqrt{r} \cos(\theta/2)\delta^2}). \quad (3.25)$$

Now, since $\theta \in [0, \pi/3)$, we have $\cos(3\theta/2), \cos(\theta/2) \in [1/\sqrt{2}, 1]$. By choosing $\delta = r^{-1/6}$, we get (3.7). \square

4 Tail bounds for random initial conditions

In this section we prove Theorem 1.1 about the upper tail decay of the $F^{(\sigma)}(s)$ distribution, which follows by combining Propositions 4.1 and 4.2 below.

Fix $c \in (0, 1)$ and let

$$\tau_c = \arg \max \left(\sqrt{2}\sigma B(t) - ct^2 \right), \quad M_c = \max_{t \in \mathbb{R}} \left(\sqrt{2}\sigma B(t) - ct^2 \right) = \sqrt{2}\sigma B(\tau_c) - c\tau_c^2 \quad (4.1)$$

be the position and the value of the maximum of the two-sided Brownian motion $\sqrt{2}\sigma B(t)$ with diffusion coefficient $2\sigma^2$ where $B(t)$ is a standard one. Note that

$$M_c = \max_{t \in \mathbb{R}} \left(\sqrt{2}\sigma B(t) - ct^2 \right) = \max_{u \in \mathbb{R}} \left(\sqrt{2}\sigma B\left(\frac{u}{2\sigma^2}\right) - \frac{cu^2}{4\sigma^4} \right) \stackrel{d}{=} \max_{u \in \mathbb{R}} \left(B(u) - \frac{c}{4\sigma^4}u^2 \right) \quad (4.2)$$

where the second equality follows by the change of variables $t = u/(2\sigma^2)$ and the third one by Brownian scaling. As a consequence, the random variable M_c has density $f_{\frac{c}{4\sigma^4}}(x)$ where f_c was defined in (3.3).

Lower bound. The idea of the lower bound is to use the inequality

$$\sup_{t \in \mathbb{R}} \left\{ \sqrt{2}\sigma B(t) + \mathcal{A}_2(t) - t^2 \right\} \geq \sqrt{2}\sigma B(t_0) + \mathcal{A}_2(t_0) - t_0^2 \quad (4.3)$$

which holds for any choice of $t_0 \in \mathbb{R}$. Furthermore, $\mathbf{P}(\sqrt{2}\sigma B(t_0) + \mathcal{A}_2(t_0) - t_0^2 > s)$ will be the largest, that is, we get the best lower bound if we take a time t_0 where $\sqrt{2}\sigma B(t_0) + \mathcal{A}_2(t_0) - t_0^2$ is the largest. As the Airy_2 process is stationary and independent of B , it does not make any difference for $\mathcal{A}_2(t_0)$ which time is chosen. Thus the idea is to choose t_0 to be the random time τ_1 which maximizes $\sqrt{2}\sigma B(t) - t^2$.

Proposition 4.1. *For all $\sigma > 0$, there is a constant C_1 independent of s such that*

$$1 - F^{(\sigma)}(s) \geq C_1 \sigma^2 (1 + 3\sigma^4)^{1/4} s^{-3/4} e^{-\frac{4}{3} \frac{1}{\sqrt{1+3\sigma^4}} s^{3/2}} \quad (4.4)$$

holds for $s \gg \max\{\sigma^{-4}, \sigma^4\}$.

Proof. The upper tail of the $F^{(\sigma)}$ distribution can be rewritten as

$$1 - F^{(\sigma)}(s) = \mathbf{E} \left(\mathbf{P} \left(\sup_{t \in \mathbb{R}} \left(\sqrt{2}\sigma B(t) + \mathcal{A}_2(t) - t^2 \right) > s \mid \tau_1 \right) \right) \quad (4.5)$$

by conditioning on the value of the time τ_1 . The conditional probability on the right-hand side of (4.5) can be lower bounded by replacing the supremum of $\sqrt{2}\sigma B(t) + \mathcal{A}_2(t) - t^2$ with its value at $t = \tau_1$ to get

$$\begin{aligned} 1 - F^{(\sigma)} &\geq \mathbf{E} \left(\mathbf{P} \left(\sqrt{2}\sigma B(\tau_1) + \mathcal{A}_2(\tau_1) - \tau_1^2 > s \mid \tau_1 \right) \right) \\ &= \mathbf{E} \left(\mathbf{P} \left(\mathcal{A}_2(\tau_1) > s - M_1 \mid \tau_1 \right) \right) \\ &= \mathbf{E} \left(1 - F_{\text{GUE}}(s - M_1) \right) \end{aligned} \quad (4.6)$$

where the first equality follows by the definition (4.1) of M_1 and by rearranging. In the second equality, we used that the Airy_2 process is stationary with GUE Tracy–Widom distribution at any position independently of τ_1 .

By using Proposition 3.1 about the asymptotic of the density of M_1 and the tail decay of the GUE Tracy–Widom distribution, see (1.7), one gets that the right-hand side of (4.6) can be lower bounded by

$$\begin{aligned} &\mathbf{E} \left(1 - F_{\text{GUE}}(s - M_1) \right) \\ &\geq \frac{4}{3} \sqrt{\frac{1}{4\sigma^4}} \frac{1}{16\pi} \int_0^s dm e^{-\frac{4}{3} \sqrt{\frac{4}{3} \frac{1}{4\sigma^4}} m^{3/2}} e^{-\frac{4}{3}(s-m)^{3/2}} \frac{\sqrt{m}}{(s-m)^{3/2}} (1 + R(s, m)) \quad (4.7) \\ &= \frac{1}{24\pi\sigma^2} \int_0^1 d\mu e^{-\frac{4}{3} \sqrt{\frac{4}{3} \frac{1}{4\sigma^4}} s^{3/2} \mu^{3/2}} e^{-\frac{4}{3} s^{3/2} (1-\mu)^{3/2}} \frac{\sqrt{\mu}}{(1-\mu)^{3/2}} (1 + R(s, s\mu)) \end{aligned}$$

with the change of variables $m = s\mu$. Here $R(s, m) = \mathcal{O}((s - m)^{-3/2}, m^{-1/4})$ and $R(s, s\mu) = \mathcal{O}(s^{-3/2}(1 - \mu)^{-3/2}, s^{-1/4}\mu^{-1/4})$ is meant as $s \rightarrow \infty$ with $\mu \in (0, 1)$.

Let $g(\mu) = -\frac{4}{3}\sqrt{\frac{4}{3}\frac{1}{4\sigma^4}\mu^{3/2} - \frac{4}{3}(1 - \mu)^{3/2}}$. One can compute that

$$g'(\mu) = 0 \text{ for } \mu = \mu_0 = \frac{3\sigma^4}{1 + 3\sigma^4} \quad (4.8)$$

as well as

$$g''(\mu) < 0 \text{ for all } \mu \in [0, 1]. \quad (4.9)$$

In particular, Taylor expansion gives $g(\mu) = g(\mu_0) - \alpha(\mu - \mu_0)^2 + \mathcal{O}((\mu - \mu_0)^3)$ with $g(\mu_0) = -\frac{4}{3}\frac{1}{\sqrt{1+3\sigma^4}}$ and $\alpha = (1 + 3\sigma^4)^{3/2}/(6\sigma^4)$.

The main contribution of the integral on the right-hand side of (4.7) comes from the regime $|\mu - \mu_0| \sim \frac{1}{s^{3/4}\sqrt{g''(\mu_0)}} = \frac{\sqrt{3}\sigma^2}{(1+3\sigma^4)^{3/4}}\frac{1}{s^{3/4}}$. We assume that $s \gg \max\{\sigma^4, \sigma^{-4}\}$ which can be written equivalently as $s^{-1/4} \ll \sigma \ll s^{1/4}$. Next we show that the error terms in R in (4.7) are small in the regime of the main contribution. If $\sigma \rightarrow 0$ as $s \rightarrow \infty$, then $\mu_0 \sim 3\sigma^4$ by (4.8) and for the regime which we consider $|\mu - \mu_0| \sim \frac{\sqrt{3}\sigma^2}{s^{3/4}} = o(\sigma^4)$ holds as long as $\sigma \gg s^{-1/4}$. Hence $\mu s \rightarrow \infty$ and $R \rightarrow 0$ in the regime of the main contribution. If $\sigma \rightarrow \infty$ with $s \rightarrow \infty$, then $1 - \mu_0 \sim \frac{1}{3\sigma^4}$ and the width of the regime considered is $\sim \frac{1}{3^{3/4}\sigma^3 s^{3/4}} = o(\sigma^{-4})$ provided that $\sigma \ll s^{1/4}$. Furthermore, $(1 - \mu)s \rightarrow \infty$ and $R \rightarrow 0$ in the regime which gives the main contribution. The error R also goes to 0 in the regime above if σ remains bounded away from 0 and infinity.

In the regime of μ that we consider, the higher order terms of the expansion are controlled by the quadratic term for all $s \gg \min\{\sigma^4, \sigma^{-4}\}$. Thus the quadratic approximation leads to the lower bound

$$\begin{aligned} \mathbf{E}(1 - F_{\text{GUE}}(s - M_1)) &\geq \frac{1}{24\sqrt{\pi}\sigma^2} \frac{\sqrt{\mu_0}}{(1 - \mu_0)^{3/2}\sqrt{\alpha}s^{3/4}} e^{-g(\mu_0)s^{3/2}} (1 + \mathcal{O}(s^{-1/4})) \\ &= \frac{\sigma^2(1 + 3\sigma^4)^{1/4}}{4\sqrt{2\pi}} s^{-3/4} e^{-\frac{4}{3}\frac{1}{\sqrt{1+3\sigma^4}}s^{3/2}} (1 + \mathcal{O}(s^{-1/4})). \end{aligned} \quad (4.10)$$

□

Upper bound. This strategy for getting the upper bound is different. We noticed that the tail distribution of $\sup_{t \in \mathbb{R}}(\mathcal{A}_2(t) - (1 - c)t^2)$ is, in the exponential scale, independent of c provided that $c < 1$. This implies that the tail distribution will be determined mostly by the tail of $M_c = \sup_{t \in \mathbb{R}}(\sqrt{2}\sigma B(t) - ct^2)$. The proof of the upper bound goes by conditioning on the value of M_c and bounding $\sqrt{2}\sigma B(t) - ct^2$ by M_c from above.

Proposition 4.2. *For all $\sigma > 0$, there is a constant C_2 independent of s such that*

$$1 - F^{(\sigma)}(s) \leq C_2\sigma^6(1 + 3\sigma^4)^{-2}s^{3/4} \ln(s) e^{-\frac{4}{3}\frac{1}{\sqrt{1+3\sigma^4}}s^{3/2}} \quad (4.11)$$

holds for $s \gg \max\{\sigma^{-4}, \sigma^4\}$.

Proof. To get an upper bound, one can write the event

$$\begin{aligned} &\left\{ \sup_{t \in \mathbb{R}} \left(\sqrt{2}\sigma B(t) + \mathcal{A}_2(t) - t^2 \right) > s \right\} \\ &= \left\{ \exists t \in \mathbb{R} : \left(\sqrt{2}\sigma B(t) - ct^2 \right) + \left(\mathcal{A}_2(t) - (1 - c)t^2 \right) > s \right\} \end{aligned} \quad (4.12)$$

for any $c \in (0, 1)$. Since the maximum of the first term on the right-hand side is M_c , it holds for any $t \in \mathbb{R}$ that $\sqrt{2}\sigma B(t) - ct^2 \leq M_c$, and one can bound the upper tail of $F^{(\sigma)}$ as

$$\begin{aligned} 1 - F^{(\sigma)}(s) &\leq \mathbf{P} \left(\exists t \in \mathbb{R} : M_c + (\mathcal{A}_2(t) - (1-c)t^2) > s \right) \\ &= \mathbf{E} \left(\mathbf{P} \left(\sup_{t \in \mathbb{R}} (\mathcal{A}_2(t) - (1-c)t^2) > s - M_c \mid M_c \right) \right) \end{aligned} \quad (4.13)$$

where the last equality follows by conditioning and rearrangement.

Now the upper bound on the right-hand side of (4.13) can be bounded by an integral using Proposition 3.1 about the density of M_c and by Theorem 1.2. Hence we get

$$\begin{aligned} 1 - F^{(\sigma)}(s) &\leq C \int_0^s dm e^{-\frac{4}{3}\sqrt{\frac{4}{3}\frac{c}{4\sigma^4}}m^{3/2}} e^{-\frac{4}{3}(s-m)^{3/2}} \frac{\sqrt{m} \ln((s-m)/(1-c))}{(s-m)^{3/4}\sqrt{1-c}} (1 + \mathcal{O}(m^{-1/4})) \\ &\leq C \int_0^1 d\mu e^{-\frac{4}{3}\sqrt{\frac{4}{3}\frac{c}{4\sigma^4}}\mu^{3/2}s^{3/2}} e^{-\frac{4}{3}(1-\mu)^{3/2}s^{3/2}} \frac{s^{3/4} \sqrt{\mu} \ln((s(1-\mu))/(1-c))}{(1-\mu)^{3/4}\sqrt{1-c}} \\ &\quad \times \left(1 + \mathcal{O}\left(\frac{1}{(s\mu)^{1/4}}\right) \right). \end{aligned} \quad (4.14)$$

As for the lower bound, we need to have $s \gg \max\{\sigma^{-4}, \sigma^4\}$ to apply the approximations.

Very similarly to (4.7), one gets that the exponent is maximal for $\mu = \mu_0 = \frac{3\sigma^4}{c+3\sigma^4}$. We also have

$$-\frac{4}{3}\sqrt{\frac{4}{3}\frac{c}{4\sigma^4}}\mu^{3/2} - \frac{4}{3}(1-\mu)^{3/2} = -\frac{4}{3}\frac{\sqrt{c}}{\sqrt{c+3\sigma^4}} - \frac{(c+3\sigma^4)^{3/2}}{6\sigma^4\sqrt{c}}(\mu - \mu_0)^2 + \mathcal{O}((\mu - \mu_0)^3). \quad (4.15)$$

This gives

$$1 - F^{(\sigma)}(s) \leq C' \frac{\sigma^4}{\sqrt{c+3\sigma^4}\sqrt{c}} \frac{\ln(s/(1-c))}{\sqrt{1-c}} e^{-\frac{4}{3}\frac{\sqrt{c}}{\sqrt{c+3\sigma^4}}s^{3/2}} (1 + \mathcal{O}(s^{-1/4})) \quad (4.16)$$

for some constant C' which does not depend on c and σ . Finally, since

$$\frac{\sqrt{c}}{\sqrt{c+3\sigma^4}} = \frac{1}{\sqrt{1+3\sigma^4}} - \frac{3\sigma^4}{2(1+3\sigma^4)^{3/2}}(1-c) + \mathcal{O}((1-c)^2), \quad (4.17)$$

we choose $1-c = \tilde{c}s^{-3/2}$. With the choice $\tilde{c} = \frac{1}{4}(1+3\sigma^4)^{3/2}/\sigma^4$, together with (4.16) we obtain

$$1 - F^{(\sigma)}(s) \leq C'' \sigma^6 (1+3\sigma^4)^{-2} s^{3/4} \ln(s) e^{-\frac{4}{3}\frac{1}{\sqrt{1+3\sigma^4}}s^{3/2}} \quad (4.18)$$

for some other constant C'' independent of σ, s . \square

5 Tail bounds for deterministic initial profile

In this section we prove Theorem 1.3 confirming the heuristics that the leading contribution for the right tail decay comes from the position where the function $h_0(t) - t^2$ is maximal.

Proof of Theorem 1.3. Let $\tau \in \mathbb{R}$ be a time such that $\kappa(h_0) = \sup_{t \in \mathbb{R}} \{h_0(t) - t^2\} = h_0(\tau) - \tau^2$. For the lower bound in (1.10) note that

$$\sup_{t \in \mathbb{R}} \{h_0(t) + \mathcal{A}_2(t) - t^2\} \geq h_0(\tau) + \mathcal{A}_2(\tau) - \tau^2 = \kappa(h_0) + \mathcal{A}_2(\tau) \quad (5.1)$$

by the definition of the time τ . Hence

$$\mathbf{P} \left(\sup_{t \in \mathbb{R}} \{h_0(t) + \mathcal{A}_2(t) - t^2\} \geq s \right) \geq \mathbf{P}(\kappa(h_0) + \mathcal{A}_2(\tau) \geq s) = 1 - F_{\text{GUE}}(s - \kappa(h_0)). \quad (5.2)$$

This inequality together with the asymptotic (1.7) leads to the lower bound in (1.10).

Now we consider the upper bound. The function $h_0(t) - t^2$ is bounded from above by $\kappa(h_0)$ for all times $t \in \mathbb{R}$ and it is bounded from above by $\kappa(h_0) - \frac{\varepsilon}{2}t^2$ for $|t| > M$. Therefore

$$\begin{aligned} & \mathbf{P} \left(\sup_{t \in \mathbb{R}} \{h_0(t) + \mathcal{A}_2(t) - t^2\} \geq s \right) \\ & \leq \mathbf{P} \left(\sup_{|t| \leq M} \{\kappa(h_0) + \mathcal{A}_2(t)\} \geq s \right) + \mathbf{P} \left(\sup_{|t| > M} \left\{ \kappa(h_0) + \mathcal{A}_2(t) - \frac{\varepsilon}{2}t^2 \right\} \geq s \right) \quad (5.3) \\ & \leq \mathbf{P} \left(\sup_{|t| \leq M} \mathcal{A}_2(t) \geq s - \kappa(h_0) \right) + \mathbf{P} \left(\sup_{t \in \mathbb{R}} \left\{ \mathcal{A}_2(t) - \frac{\varepsilon}{2}t^2 \right\} \geq s - \kappa(h_0) \right). \end{aligned}$$

The first term is bounded using Lemma 2.1. The second term is bounded using Theorem 1.2. Altogether we get

$$\begin{aligned} & \mathbf{P} \left(\sup_{t \in \mathbb{R}} \{h_0(t) + \mathcal{A}_2(t) - t^2\} \geq s \right) \\ & \leq C' \frac{2M}{(s - \kappa(h_0))^{1/4}} e^{-\frac{4}{3}(s - \kappa(h_0))^{3/2}} + C \frac{\ln[2(s - \kappa(h_0))/\varepsilon]}{(s - \kappa(h_0))^{3/4} \sqrt{\varepsilon/2}} e^{-\frac{4}{3}(s - \kappa(h_0))^{3/2}}. \quad (5.4) \end{aligned}$$

Since ε is fixed, for large s the second term is smaller than the first one, which completes the proof. \square

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