

# Exercises in Extreme value theory

2022 spring semester

1. Show that  $L(t) = \log t$  is a slowly varying function but  $t^\epsilon$  is not if  $\epsilon \neq 0$ .
2. **Homework 1.A (2nd Mar 2022)** If the random variable  $X$  has distribution  $F$  with finite variance, then show that the condition

$$\lim_{x \rightarrow \infty} \frac{x^2 \mathbf{P}(|X| > x)}{\int_{|y| \leq x} y^2 dF(y)} = 0$$

automatically holds.

*Hint:* Prove and use the inequality

$$x^2 \mathbf{P}(|X| > x) \leq \int_{|y| \geq x} y^2 dF(y).$$

3. Let  $W_1$  and  $W_2$  be independent standard normal random variables.
  - (a) Check that  $1/W_2^2$  has the Lévy distribution, that is, the stable law with index  $\alpha = 1/2$  and  $\kappa = 1$ , i.e. it has density

$$\frac{1}{\sqrt{2\pi y^3}} \exp\left(-\frac{1}{2y}\right) \quad \text{for } y > 0.$$

- (b) Prove that  $W_1/W_2$  has Cauchy distribution.

4. **Homework 1.B (2nd Mar 2022)** Let  $\tau = \inf\{t > 0 : B_t = 1\}$  be the hitting time of level 1 by the standard Brownian motion  $B_t$  which starts from  $B_0 = 0$ . Show that  $\tau$  has the Lévy distribution, that is, its density is  $(2\pi y^3)^{-1/2} \exp(-1/(2y))$  for  $y > 0$ .

*Hint:* It can be used without proof that as a consequence of the reflection principle

$$\mathbf{P}(\max\{B_s : s \in [0, t]\} > x) = 2 \left(1 - \Phi\left(\frac{x}{\sqrt{t}}\right)\right)$$

where  $\Phi(x) = \int_{-\infty}^x (2\pi)^{-1/2} \exp(-y^2/2) dy$  is the standard normal distribution function. The distribution function of  $\tau$  can be directly expressed by the probability in the equality above.

5. **Homework 1.C (2nd Mar 2022)** Let  $X$  have a symmetric stable law with index  $\alpha$ . Show that  $\mathbf{E}|X|^p < \infty$  for  $p \in (0, \alpha)$ .

*Hint:* If  $\varphi(t) = \mathbf{E}(e^{itX})$  denotes the characteristic function of  $X$ , then the following inequality can be used without proof for  $u > 0$ :

$$\mathbf{P}\left(|X| > \frac{2}{u}\right) \leq \frac{1}{u} \int_{-u}^u (1 - \varphi(t)) dt.$$

Further, the identity

$$\mathbf{E}(|X|^p) = p \int_0^\infty t^{p-1} \mathbf{P}(|X| > t) dt$$

holds.

6. Let  $X_1, X_2, \dots$  be an iid. sequence of Poisson random variables with parameter  $\lambda > 0$ , that is,

$$\mathbf{P}(X = k) = e^{-\lambda} \frac{\lambda^k}{k!}$$

for  $k = 0, 1, 2, \dots$  and let  $M_n := \max_{1 \leq i \leq n} X_i$ . Show that there is no such normalization under which the sequence of maxima  $M_n$  has a non-degenerate limit law.

*Hint:* Show that the necessary condition

$$\lim_{x \uparrow x_F} \frac{\overline{F}(x+)}{\overline{F}(x)} = 1$$

fails to hold along a sequence of integers where  $\overline{F}(x) = 1 - F(x)$  is the tail probability function. More precisely  $\overline{F}(k+1)/\overline{F}(k)$  converge to 0 for the integers  $k \rightarrow \infty$ .

7. **Homework 2.A (23rd Mar 2022)** Let  $X_1, X_2, \dots$  be an iid. sequence of negative binomial random variables with parameters  $p \in (0, 1)$  and  $m \in \mathbb{N}$ , that is,

$$\mathbf{P}(X = k) = \binom{k+m-1}{k} p^m (1-p)^k$$

for  $k = 0, 1, 2, \dots$  and let  $M_n := \max_{1 \leq i \leq n} X_i$ . Show that there is no such normalization under which the sequence of maxima  $M_n$  has a non-degenerate limit law.

*Hint:* Show that the necessary condition

$$\lim_{x \uparrow x_F} \frac{\overline{F}(x+)}{\overline{F}(x)} = 1$$

fails to hold along a sequence of integers where  $\overline{F}(x) = 1 - F(x)$  is the tail probability function. More precisely  $\overline{F}(k+1)/\overline{F}(k)$  converges to  $1-p$  for the integers  $k \rightarrow \infty$ . To this end, show and use the asymptotic identity

$$\mathbf{P}(X = k) = \frac{k^{m-1}}{(m-1)!} p^m (1-p)^k (1 + o(1))$$

as  $k \rightarrow \infty$ .

8. **Homework 2.B (23rd Mar 2022)** Suppose that a function  $f : \mathbb{R} \rightarrow \mathbb{R}$  satisfies Cauchy's functional equation

$$f(x+y) = f(x) + f(y)$$

for all  $x, y \in \mathbb{R}$ . Prove that the value of the function  $f(z)$  at any  $z \in \mathbb{R}$  determines all the values  $f\left(\frac{p}{q}z\right)$  for any  $\frac{p}{q} \in \mathbb{Q}$  rational.

9. For  $f : \mathbb{R} \rightarrow \mathbb{R}$  functions, consider Cauchy's functional equation

$$f(x+y) = f(x) + f(y).$$

Prove that the only solutions for the equation are the linear functions  $f(x) = \alpha x$  for some  $\alpha \in \mathbb{R}$ , if we assume that  $f$  is

- (a) continuous;
- (b) continuous at one point;
- (c) monotonic on any interval;
- (d) **Homework 2.C (23rd Mar 2022)** bounded on any interval;
- (e) measurable.

*Hint:* Lusin's theorem ensures that by the measurability of  $f$  there is a subset  $F \subseteq [0, 1]$  with Lebesgue measure  $2/3$  such that  $f$  is uniformly continuous on  $F$ . Then for any  $h \in (0, 1/3)$ , the sets  $F$  and  $F - h = \{x - h : x \in F\}$  cannot be disjoint. Use this to conclude that the function  $f$  is bounded on an interval  $(0, \delta)$  for some small  $\delta > 0$ .

10. Let  $X_1, X_2, \dots$  be random variables defined on the probability space  $(\Omega, \mathcal{A}, \mathbf{P})$  which have the same distribution function  $F(x)$ , but we do not assume anything about their dependence structure. Let  $M_n := \max_{1 \leq i \leq n} |X_i|$ .

- (a) Suppose that there is an  $\alpha > 0$  such that  $\int_{-\infty}^{\infty} |x|^\alpha dF(x) < \infty$ , that is,  $\mathbf{E}(|X_i|^\alpha) < \infty$ . Prove that for any  $\varepsilon > 0$  and for fixed  $\delta > 0$ ,

$$\lim_{n \rightarrow \infty} \mathbf{P} \left( n^{-(1/\alpha + \varepsilon)} |M_n| > \delta \right) = 0.$$

- (b) Suppose that there is an  $s > 0$  such that  $\int_{-\infty}^{\infty} \exp(s|x|) dF(x) < \infty$ , that is,  $\mathbf{E}(\exp(s|X_i|)) < \infty$ . Prove that for any sequence  $b_n$  that goes to  $\infty$  and for fixed  $\delta > 0$ ,

$$\lim_{n \rightarrow \infty} \mathbf{P} \left( (b_n \log n)^{-1} |M_n| > \delta \right) = 0.$$

*Hint:* Use the following inequality

$$\mathbf{P} \left( \max_{1 \leq i \leq n} |X_i| > \lambda \right) = \mathbf{P}(\cup_{i=1}^n \{|X_i| > \lambda\}) \leq \sum_{i=1}^n \mathbf{P}(|X_i| > \lambda) = n\mathbf{P}(|X_1| > \lambda)$$

and a Markov type inequality.

11. **Homework 3.A (13th Apr 2022)** Let  $X_1, X_2, \dots$  be independent and identically distributed random variables with common distribution function  $F(x) = \mathbf{P}(X_i < x)$ . Let  $M_n := \max_{1 \leq i \leq n} X_i$ . If  $F(x) < 1$  for all  $x < \infty$  and  $\lim_{x \rightarrow \infty} x^\alpha (1 - F(x)) = b$  for some fixed constants  $\alpha, b \in (0, \infty)$  (that is,  $1 - F(x) \sim bx^{-\alpha}$  as  $x \rightarrow \infty$ ), then show that the distribution of  $(bn)^{-1/\alpha} M_n$  converges weakly to the Fréchet distribution:

$$\mathbf{P} \left( (bn)^{-1/\alpha} M_n < x \right) \rightarrow \mathbb{1}(x > 0) \exp(-x^{-\alpha}).$$

12. Let  $X_1, X_2, \dots$  be independent and identically distributed random variables with common distribution function  $F(x) = \mathbf{P}(X_i < x)$ . Let  $M_n := \max_{1 \leq i \leq n} X_i$ . If  $F(x) < 1$  for all  $x < \infty$  and  $\lim_{x \rightarrow \infty} e^{\lambda x} (1 - F(x)) = b$  for some fixed constants  $\lambda, b \in (0, \infty)$  (that is,  $1 - F(x) \sim be^{-\lambda x}$  as  $x \rightarrow \infty$ ), then show that the distribution of  $\lambda M_n - \log(bn)$  converges weakly to the Gumbel distribution:

$$\mathbf{P} \left( \lambda M_n - \log(bn) < x \right) \rightarrow \exp(-e^{-x}).$$

13. Let  $X_1, X_2, \dots$  be independent and identically distributed random variables with common distribution function  $F(x) = \mathbf{P}(X_i < x)$ . Let  $M_n := \max_{1 \leq i \leq n} X_i$ . If  $F(x_0) = 1$  and  $F(x) < 1$  for all  $x < x_0$  and  $\lim_{x \rightarrow x_0} (x_0 - x)^{-\alpha} (1 - F(x)) = b$  for some fixed constants  $\alpha, b \in (0, \infty)$  (that is,  $1 - F(x) \sim b(x_0 - x)^\alpha$  as  $x \rightarrow x_0$ ), then show that the distribution of  $(bn)^{1/\alpha} (M_n - x_0)$  converges weakly to the Weibull distribution:

$$\mathbf{P}((bn)^{1/\alpha} (M_n - x_0) < x) \rightarrow \mathbb{1}(x < 0) \exp(-(-x)^\alpha) + \mathbb{1}(x \geq 0).$$

14. **Homework 3.B (13th Apr 2022)** Let  $X_1, X_2, \dots$  be independent and identically distributed random variables with common distribution function  $F(x) = \mathbf{P}(X_i < x)$ . Suppose that the common distribution has a finite right endpoint, that is,

$$x_F = \sup\{x \in \mathbb{R} : F(x) < 1\} < \infty$$

and that the right endpoint has a positive mass  $\mathbf{P}(X_i = x_F) = \overline{F}(x_F) > 0$ . Let  $M_n := \max_{1 \leq i \leq n} X_i$ . Prove that for any sequence of reals  $(u_n)$ , if the sequence of probabilities  $\mathbf{P}(M_n < u_n)$  converges, then the limit is either 0 or 1.

*Hint:* Use the proposition about Poisson approximation to see that it is enough to show that  $n\overline{F}(u_n)$  goes either to 0 or to  $\infty$ .

15. **Homework 3.C (13th Apr 2022)** Let  $X$  be a random variable with distribution function  $F$  where the tail is  $1 - F(x) = x^{-\alpha}$  for  $x \geq 1$  with some  $\alpha > 0$ . Then we know that  $F \in \text{MDA}(\Phi_\alpha)$ . Which MDA does the distribution of  $X^p$  and that of  $\ln(X)$  belong to if  $p > 0$ ? What are the normalization constants?
16. Suppose that  $X > 0$  is a random variable and  $\alpha > 0$  is a number. Show that the following are equivalent:

- (a)  $X$  has the Fréchet distribution function

$$\Phi_\alpha(x) = \begin{cases} 0 & \text{if } x \leq 0 \\ \exp(-x^{-\alpha}) & \text{if } x > 0 \end{cases}$$

- (b)  $\ln X^\alpha$  has the Gumbel distribution function

$$\Lambda(x) = \exp(-e^{-x})$$

- (c)  $-X^{-1}$  has the Weibull distribution function

$$\Psi_\alpha(x) = \begin{cases} \exp(-(-x)^\alpha) & \text{if } x \leq 0 \\ 1 & \text{if } x > 0 \end{cases}$$

Show the equivalence of (a) and (b).

17. Let  $X_1, X_2, \dots$  be an independent but not identically distributed sequence of random variables, let  $X_k$  be exponential with parameter  $\lambda_k$ . Denote by  $m_n = \min(X_1, \dots, X_n)$  the minimum record up to  $n$  and let  $m_\infty = \lim_{n \rightarrow \infty} m_n$  which exists since  $m_n$  is non-increasing.

- (a) Show that if  $\sum_{k=1}^{\infty} \lambda_k < \infty$ , then with probability one, the minimum record is broken finitely many times and  $m_\infty > 0$ .
- (b) Show that if  $\sum_{k=1}^{\infty} \lambda_k = \infty$ , then with probability one, the minimum record is broken infinitely many times and  $m_\infty = 0$ .

*Hint:* One can use the representation of  $X_k$  as the first point of a Poisson point process on  $\mathbb{R}_+$  with intensity  $\lambda_k$  and the fact that the union of independent Poisson point processes with intensities  $\lambda_k$  for  $k = 1, 2, \dots$  is a Poisson point process with intensity  $\sum_{k=1}^{\infty} \lambda_k$ .

18. Let  $X_1, X_2, \dots$  be an independent but not identically distributed sequence of random variables, let  $X_k$  be exponential with parameter  $\lambda_k$ . Denote by  $M_n = \max(X_1, \dots, X_n)$  the maximum record up to  $n$  and let  $M_\infty = \lim_{n \rightarrow \infty} M_n$  which exists since  $M_n$  is non-decreasing.

- (a) Show that for any constant  $K > 0$  if

$$\sum_{n=1}^{\infty} e^{-\lambda_n K} = \infty,$$

then with probability one there are infinitely many indices  $n$  such that  $X_n > K$ . Further, for any  $K > 0$  if

$$\sum_{n=1}^{\infty} e^{-\lambda_n K} < \infty,$$

then with probability one there are finitely many indices  $n$  such that  $X_n > K$ .

*Hint:* Use the Borel–Cantelli lemmas.

- (b) Conclude that if  $\sum_{n=1}^{\infty} e^{-\lambda_n K} = \infty$  for any  $K > 0$ , then with probability one,  $M_\infty = \infty$  and the maximum record is broken infinitely many times. If  $\sum_{n=1}^{\infty} e^{-\lambda_n K} < \infty$  for any  $K > 0$ , then with probability one,  $M_\infty < \infty$  and the maximum record is broken finitely many times.
- (c) If there are  $K_1 > K_2 > 0$  such that  $\sum_{n=1}^{\infty} e^{-\lambda_n K_1} < \infty$  but  $\sum_{n=1}^{\infty} e^{-\lambda_n K_2} = \infty$ , then there is a critical  $K_c$  such that for smaller values of  $K$  the sum is infinite and for larger values of  $K$  the sum is finite. Show that in this case with probability one

$$\limsup_{n \rightarrow \infty} X_n = K_c.$$

- (d) Suppose that there is a critical  $K_c$  as described above. Show that if

$$\sum_{n=1}^{\infty} e^{-\lambda_n K_c} = \infty,$$

then with probability one the maximum record is broken finitely many times.

*Hint:* The condition shows that there are infinitely many indices  $n$  such that  $X_n > K_c$ , but for any  $\varepsilon > 0$ , there are only finitely many with  $X_n > K_c + \varepsilon$ .

- (e) Suppose that there is a critical  $K_c$  as described above. Show that if

$$\sum_{n=1}^{\infty} e^{-\lambda_n K_c} < \infty,$$

then with positive probability the maximum record is broken infinitely many times. Compute this probability.

*Hint:* Observe that the maximum record is broken infinitely many times if and only if  $X_n < K_c$  for all indices  $n$ , since  $\limsup_{n \rightarrow \infty} X_n = K_c$ .