

Ex: $X \sim N(m, \sigma^2)$ NORMAL DISTRIBUTION

DENSITY: $\varphi_{m, \sigma}(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-m)^2}{2\sigma^2}}$

MOMENT GENERATING FUNCTION OF X ?

FIRST: $Y \sim N(0, 1)$

$$\begin{aligned} Z_Y(\lambda) &= E(e^{\lambda Y}) = \int_{-\infty}^{\infty} e^{\lambda x} \cdot \varphi_{0,1}(x) dx = \\ &\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left(\lambda x - \frac{x^2}{2}\right) dx \stackrel{?}{=} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(x-\lambda)^2}{2} + \frac{\lambda^2}{2}\right) dx \\ &= e^{\lambda^2/2} \int_{-\infty}^{\infty} \varphi_{\lambda,1}(x) dx = e^{\lambda^2/2} \quad \text{NOW } X = m + \sigma \cdot Y \stackrel{?}{=} \end{aligned}$$

$$\begin{aligned} Z_X(\lambda) &= E(e^{\lambda X}) = E\left(e^{\lambda \cdot m + \lambda \cdot \sigma Y}\right) = \\ &e^{\lambda \cdot m} \cdot E\left(e^{\lambda \cdot \sigma Y}\right) = e^{\lambda \cdot m} \cdot Z_Y(\lambda \cdot \sigma) = e^{\lambda \cdot m + \frac{\sigma^2 \lambda^2}{2}} \end{aligned}$$

LOG. MOM GEN. F'N OF X : $\hat{I}(\lambda) = \lambda \cdot m + \frac{\sigma^2 \lambda^2}{2}$

LEGENDRE TRANSFORM OF \hat{I} :

$$I(x) = \max_{\lambda} \{ \lambda \cdot x - \hat{I}(\lambda) \}$$

$$\frac{d}{d\lambda} (\lambda \cdot x - \hat{I}(\lambda)) = x - m - \lambda \cdot \sigma^2 = 0 \Rightarrow$$

$$\lambda^* = \frac{x-m}{\sigma^2}$$

$$\begin{aligned} I(x) &= \lambda^* \cdot x - \hat{I}(\lambda^*) = \frac{x-m}{\sigma^2} \cdot x - \frac{x-m}{\sigma^2} \cdot m - \frac{\sigma^2 \cdot \left(\frac{x-m}{\sigma^2}\right)^2}{2} = \\ &= \frac{(x-m)^2}{2\sigma^2} \end{aligned}$$

THUS IF X_1, X_2, \dots I.I.D. $\mathcal{N}(\mu, \sigma^2)$

$S_n = X_1 + \dots + X_n$ AND $X > m$ THEN

$$P\left(\frac{S_n}{n} \geq x\right) \leq \exp(-n \cdot I(x)) = \exp\left(-n \cdot \frac{(x-\mu)^2}{2\sigma^2}\right)$$

(SEE PAGE 8, 9)

IS THIS UPPER BOUND SHARP?

NOTE: $S_n \sim \mathcal{N}(n\mu, \sigma^2 \sqrt{n})$ $\frac{S_n}{n} \sim \mathcal{N}(\mu, \frac{\sigma^2}{\sqrt{n}})$

$$\text{THUS } P\left(\frac{S_n}{n} \geq x\right) = \int_x^\infty \frac{\sqrt{n}}{\sqrt{2\pi} \cdot \sigma} \cdot \exp\left(-n \cdot \frac{(y-\mu)^2}{2\sigma^2}\right) dy$$

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log\left(P\left(\frac{S_n}{n} \geq x\right)\right) =$$

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log\left(\frac{\sqrt{n}}{\sqrt{2\pi} \cdot \sigma}\right) + \lim_{n \rightarrow \infty} \frac{1}{n} \log\left(\int_x^\infty e^{-n \cdot I(y)} dy\right)$$

$\Rightarrow \lim_{n \rightarrow \infty} -\inf_{y \geq x} I(y) = I(x)$

$$\text{THUS INDEED: } P\left(\frac{S_n}{n} \geq x\right) \approx e^{-n \cdot I(x)}$$

SO THE UPPER BOUND (SEE PAGE 2)
FOR NOTATION
WAS SHARP IN THIS SENSE

$Z(\lambda) := E(e^{\lambda \cdot X})$, ASSUME THAT $\exists \underline{\lambda} < 0, \bar{\lambda} > 0$

SUCH THAT $Z(\lambda) < +\infty$ IF $\lambda \in [\underline{\lambda}, \bar{\lambda}]$ ■

LEMMA: $\lambda \mapsto Z(\lambda)$ IS ANALYTIC ■ ON $(\underline{\lambda}, \bar{\lambda})$,

moreover $\frac{d^n}{d\lambda^n} Z(\lambda) = \mathbb{E}(X^n \cdot e^{\lambda \cdot X})$

PROOF: $\frac{d^n}{d\lambda^n} e^{\lambda \cdot X} = X^n \cdot e^{\lambda \cdot X}$, LET $\lambda_0 \in (\underline{\lambda}, \bar{\lambda})$

THUS $\exists \varepsilon > 0 : [\lambda_0 - \varepsilon, \lambda_0 + \varepsilon] \subseteq (\underline{\lambda}, \bar{\lambda})$

$$e^{(\lambda_0 + M) \cdot X} = e^{\lambda_0 \cdot X} \cdot e^{M \cdot X} = \sum_{n=0}^{\infty} e^{\lambda_0 \cdot X} \cdot \frac{M^n \cdot X^n}{n!} \quad \boxed{\text{IF } |M| \leq \varepsilon}$$

NOTE: $\sum_{n=0}^{\infty} \frac{|M|^n \cdot |X|^n}{n!} = e^{|M| \cdot |X|} \leq e^{\varepsilon \cdot |X|} \leq e^{-\varepsilon \cdot X} + e^{\varepsilon \cdot X}$

THUS $\forall n \geq 0 : \underbrace{\sum_{n=0}^{\infty} \left| e^{\lambda_0 \cdot X} \cdot \frac{M^n \cdot X^n}{n!} \right|}_{\text{U}} \leq \underbrace{e^{(\lambda_0 - \varepsilon) \cdot X} + e^{(\lambda_0 + \varepsilon) \cdot X}}$

THUS $E(-||-||) \leq E(-||-) < +\infty$ ■

THUS BY DOMINATED CONVERGENCE THM:

$$E(e^{(\lambda_0 + M) \cdot X}) = E\left(\lim_{n \rightarrow \infty} \sum_{n=0}^{\infty} e^{\lambda_0 \cdot X} \cdot \frac{M^n \cdot X^n}{n!}\right) = \boxed{\text{ }} \quad \boxed{\text{ }}$$

$$= \lim_{n \rightarrow \infty} E\left(\sum_{n=0}^{\infty} e^{\lambda_0 \cdot X} \cdot \frac{M^n \cdot X^n}{n!}\right) = \lim_{n \rightarrow \infty} \sum_{n=0}^{\infty} \frac{M^n}{n!} E(X^n \cdot e^{\lambda_0 \cdot X})$$

$$= \sum_{n=0}^{\infty} \frac{M^n}{n!} E(X^n \cdot e^{\lambda_0 \cdot X}) \quad \boxed{\text{ }} \quad \text{IF } |M| \leq \varepsilon \quad \checkmark$$

$$\hat{I}(\lambda) = \ln(Z(\lambda))$$

WANT: $\lambda \mapsto \hat{I}(\lambda)$ IS STRICTLY CONVEX

$$\hat{I}'(\lambda) = \frac{Z'(\lambda)}{Z(\lambda)} = \frac{\mathbb{E}(X \cdot e^{\lambda X})}{Z(\lambda)}$$

$$\hat{I}''(\lambda) = \frac{Z''(\lambda) \cdot Z(\lambda) - (Z'(\lambda))^2}{(Z(\lambda))^2} = \frac{\mathbb{E}(X^2 \cdot e^{\lambda X})}{Z(\lambda)} - (\hat{I}'(\lambda))^2$$

WHY DO WE HAVE $\boxed{\hat{I}''(\lambda) > 0}$?

NOTE: $Z(0) = 1$, $\hat{I}(0) = 0$, $\hat{I}'(0) = \mathbb{E}(X)$

$$\hat{I}''(0) = \mathbb{E}(X^2) - (\mathbb{E}(X))^2 = \text{Var}(X) > 0$$

AND FOR $\lambda \neq 0$, $\hat{I}''(\lambda) = \text{Var}(\text{?})$

DEF: (MEASURE CHANGE, EXPONENTIALLY TILTED R.V.'S)

GIVEN X AND $\lambda \in \mathbb{R}$ SUCH THAT $Z(\lambda) < +\infty$,

DEFINE $X^{(\lambda)}$ TO BE A R.V. THAT SATISFIES

$$P(X^{(\lambda)} \in A) := \mathbb{E}\left(\frac{e^{\lambda X} \cdot \mathbb{I}[X \in A]}{Z(\lambda)}\right) \quad \text{FOR ANY } A \subseteq \mathbb{R}$$

IN PARTICULAR:

$$F^{(\lambda)}(x) := P(X^{(\lambda)} \leq x) := \mathbb{E}\left(\frac{e^{\lambda X} \cdot \mathbb{I}[X \leq x]}{Z(\lambda)}\right)$$

NOTE: $F^{(\lambda)}$ IS INDEED A PROB. DISTRIBUTION F_N :
NON-DECREASING, RIGHT-CONTINUOUS,

$$\lim_{x \rightarrow -\infty} F^{(\lambda)}(x) = 0 \quad \text{MONOTONE CONVERGENCE THM}$$

$$\lim_{x \rightarrow +\infty} F^{(\lambda)}(x) = \mathbb{E}\left(\frac{e^{\lambda X}}{Z(\lambda)} \cdot \mathbb{I}[X < +\infty]\right) = \frac{z(\lambda)}{Z(\lambda)} = 1$$

NOTE: $\mathbb{E}(g(X^{(\lambda)})) \stackrel{c}{=} \mathbb{E}\left(g(X) \cdot \frac{e^{\lambda X}}{Z(\lambda)}\right)$

NOTE: IF X IS INTEGER-VALUED WITH

$$P(X=r) = p_r$$

$$r \in \mathbb{Z}$$

THEN

$$P(X^{(\lambda)}=r) = \frac{e^{\lambda r} \cdot p_r}{Z(\lambda)}$$

NOTE: IF X IS ABSOLUTELY CONTINUOUS WITH
DENSITY FUNCTION $f(x)$ THEN

$$P(X^{(\lambda)} \leq x) = \int_{-\infty}^x \frac{e^{\lambda y}}{Z(\lambda)} \cdot f(y) dy, \text{ THUS } X^{(\lambda)} \text{ IS}$$

ALSO ABS. CONTINUOUS WITH DENSITY

$$\text{FUNCTION } f^{(\lambda)}(x) \stackrel{H}{=} \frac{e^{\lambda x}}{Z(\lambda)} \cdot f(x)$$

NOW $Z'(\lambda) = \mathbb{E}(X \cdot e^{\lambda X})$, $Z''(\lambda) = \mathbb{E}(X^2 \cdot e^{\lambda X})$,

$$\hat{I}'(\lambda) = \frac{Z'(\lambda)}{Z(\lambda)} = \mathbb{E}\left(X \cdot \frac{e^{\lambda X}}{Z(\lambda)}\right) \stackrel{H}{=} \mathbb{E}(X^{(\lambda)})$$

THUS:

$$\hat{I}''(\lambda) \stackrel{\text{E}}{=} \frac{\hat{z}''(\lambda)}{\hat{z}(\lambda)} - \left(\frac{\hat{z}'(\lambda)}{\hat{z}(\lambda)} \right)^2 \quad \blacksquare = \mathbb{E}((\hat{X}^{(\lambda)})^2) - (\mathbb{E}(\hat{X}^{(\lambda)}))^2$$

$$\boxed{\hat{I}''(\lambda) \stackrel{\text{W}}{=} \text{Var}(\hat{X}^{(\lambda)}) > 0}$$

INDEED:

\hat{I} STRICTLY CONVEX!

THUS BY PAGE 12-13:

$$\text{IF WE DEFINE } I(x) = \sup_{\lambda} \{ \lambda \cdot x - \hat{I}(\lambda) \}$$

$$\text{THEN INDEED } \hat{I}(\lambda) = \sup_x \{ \lambda \cdot x - I(x) \}, \quad \blacksquare$$

MOREOVER

$\lambda \mapsto \hat{I}'(\lambda)$ IS STRICTLY INCREASING, CONTINUOUS

$$x \mapsto I'(x) \text{ IS } -/-/-/-/-/-/-$$

$$\boxed{I' = (\hat{I}')^{-1}}$$

AND IF WE DEFINE

$$\boxed{\lambda^*(x) := (\hat{I}')^{-1}(x)}$$

$$\text{THEN } \mathbb{E}(\hat{X}^{(\lambda^*(x))}) = \hat{I}'(\lambda^*(x)) \quad \blacksquare = x \text{ AND}$$

$$\boxed{\hat{I}(x) = x \cdot \lambda^*(x) - \hat{I}(\lambda^*(x))}$$

$$\boxed{\hat{I}''(x) = \frac{1}{\text{Var}(\hat{X}^{(\lambda^*(x))})}}$$

NOTE: $\hat{I}(0) = 0$, $\hat{I}'(0) = \mathbb{E}(\hat{X}) = m$, THUS

$$\lambda^*(m) = 0, \text{ THUS } I(m) = x \cdot 0 - \hat{I}(0) = 0$$

($I'(m)$) \blacksquare