

RECALL: CHARACTERISTIC FUNCTION OF THE RANDOM VARIABLE \mathbb{X} :

$$\varphi: \mathbb{R} \rightarrow \mathbb{C} \quad \varphi(t) := \mathbb{E}(e^{it \cdot \mathbb{X}})$$

PROPERTIES: A $\varphi(0) = 1$ ✓

B $|\varphi(t)| \leq 1$ C PROOF: $|\mathbb{E}(e^{it \cdot \mathbb{X}})| \leq \mathbb{E}(|e^{it \cdot \mathbb{X}}|) = 1$

D $\varphi(-t) = \overline{\varphi(t)}$ E PROOF: $\varphi(-t) = \mathbb{E}(\cos(-t \cdot \mathbb{X})) + i \cdot \mathbb{E}(\sin(-t \cdot \mathbb{X})) = \overline{\varphi(t)}$

$$\mathbb{E}(\cos(t \cdot \mathbb{X})) - i \cdot \mathbb{E}(\sin(t \cdot \mathbb{X})) = \overline{\varphi(t)}$$

F F $t \mapsto \varphi(t)$ IS CONTINUOUS G PROOF: $\lim_{t \rightarrow t_0} \varphi(t) =$

$$\lim_{t \rightarrow t_0} \mathbb{E}(e^{it \cdot \mathbb{X}}) \stackrel{\text{G}}{=} \mathbb{E}\left(\lim_{t \rightarrow t_0} e^{it \cdot \mathbb{X}}\right) = \varphi(t_0)$$

BY DOMINATED CONV. THM. (SEE PAGE 37):

$|e^{it \cdot \mathbb{X}}| \leq 1$ FOR ALL t , THUS 1 IS THE DOM. R.V.

IF \mathbb{X} IS AN INTEGER-VALUED R.V. THEN:

H $\varphi(t+2\pi) = \varphi(t)$ I PROOF: $\varphi(t+2\pi) =$

$$\mathbb{E}(e^{it \cdot \mathbb{X}} \cdot e^{i2\pi \mathbb{X}}) \stackrel{\text{J}}{=} \mathbb{E}(e^{it \cdot \mathbb{X}}) = \varphi(t). \text{ IN PARTICULAR:}$$

K $\varphi(n \cdot 2\pi) = \varphi(0) = 1, n \in \mathbb{Z}$

$t \mapsto \varphi(t)$ IS OF "POSITIVE TYPE":

IF $t_1, t_2, \dots, t_n \in \mathbb{R}$ AND IF $\underline{M} \in \mathbb{C}^{n \times n}$

$M(k, l) := \varphi(t_k - t_l)$, $1 \leq k, l \leq n$, THEN

\underline{M} IS A POSITIVE SEMIDEFINITE HERMITIAN MATRIX.

PROOF: WANT: $\underline{z}^T \underline{M} \underline{z} \geq 0$ $\forall \underline{z} = (z_1, \dots, z_n) \in \mathbb{C}^n$

$$\underline{z}^T \underline{M} \underline{z} = \sum_{k, l=1}^n z_k \cdot \varphi(t_k - t_l) \cdot \bar{z}_l =$$

$$= \sum_{k, l=1}^n z_k \cdot \mathbb{E}(e^{i \cdot t_k \cdot X} \cdot e^{-i \cdot t_l \cdot X}) \cdot \bar{z}_l =$$
$$= \mathbb{E}\left(\sum_{k, l=1}^n z_k \cdot e^{i \cdot t_k \cdot X} \cdot \overline{(e^{i \cdot t_l \cdot X} \cdot z_l)}\right)$$

$$= \mathbb{E}\left(\left|\sum_{k=1}^n z_k \cdot e^{i \cdot t_k \cdot X}\right|^2\right) \geq 0 \quad \checkmark$$

G THM (BOCHNER): IF $\varphi: \mathbb{R} \rightarrow \mathbb{C}$ IS

CONTINUOUS AT $t=0$, $\varphi(0)=1$ AND φ

IS OF POSITIVE TYPE THEN THERE IS
A R.V. X SUCH THAT $\varphi(t) = \mathbb{E}(e^{i \cdot t \cdot X})$.

(WE WILL NOT PROVE THIS AND WE WILL
NOT USE THIS.)

IF $Y = a \cdot X + b$ THEN $\varphi_Y(t) = e^{i \cdot t \cdot b} \cdot \varphi_X(at)$

PROOF:

$$\varphi_Y(t) = \mathbb{E}(e^{i \cdot t \cdot (aX + b)}) \stackrel{\text{B}}{=} e^{i \cdot t \cdot b} \cdot \underbrace{\mathbb{E}(e^{i \cdot a \cdot t \cdot X})}_{\varphi_X(at)}$$

IF X, Y ARE INDEP., $Z = X + Y$, THEN

$\varphi_Z(t) \stackrel{\text{C}}{=} \varphi_X(t) \cdot \varphi_Y(t)$

PROOF:

$$\varphi_Z(t) = \mathbb{E}(e^{i \cdot t \cdot X} \cdot e^{i \cdot t \cdot Y}) \stackrel{\text{D}}{=} \mathbb{E}(e^{i \cdot t \cdot X}) \cdot \mathbb{E}(e^{i \cdot t \cdot Y})$$

INDEPENDENCE

$\varphi_X(t)$ $\varphi_Y(t)$

EX: $Y \sim \text{UNI}[-a, a]$, THEN $\varphi_Y(t) = ?$

NOTE: $Y = a \cdot X$, $X \sim \text{UNI}[-1, 1]$ E

NOTE: $X \stackrel{\text{F}}{\sim} (-X)$, SO $\varphi_X(t) \stackrel{\text{G}}{=} \mathbb{E}(\cos(tX)) \stackrel{\text{H}}{=}$

$$= \int_{-1}^1 \cos(tx) \cdot \frac{1}{2} dx = \frac{1}{2} \left[\frac{1}{t} \sin(tx) \right]_{-1}^1 = \frac{2 \cdot \sin(t)}{2t}$$

THUS $\varphi_Y(t) = \frac{\sin(at)}{at}$

THM: IF $E(|X|^r) < +\infty$ THEN φ_X IS
r TIMES DIFFERENTIABLE AND

$$\varphi_X^{(r)}(t) = \underset{A}{E} \left[(i \cdot X)^r \cdot e^{i \cdot t \cdot X} \right] \quad \square \quad \square \quad \square$$

IN PARTICULAR:
$$\boxed{\varphi_X^{(r)}(0) = \underset{B}{i^r \cdot E(X^r)}} \quad \square$$

PROOF: FORMALLY SIMPLE:

$$\frac{d^r}{dt^r} E(e^{i \cdot t \cdot X}) \stackrel{C}{=} E \left(\frac{d^r}{dt^r} e^{i \cdot t \cdot X} \right) \stackrel{D}{=} E \left[(i \cdot X)^r \cdot e^{i \cdot t \cdot X} \right]$$

BUT CAN YOU CHANGE LIMITS WITH E?

YES, BUT WE WILL ONLY GIVE A PROOF
IN THE CASE WHEN $r=1$ AND $t=0$. \square

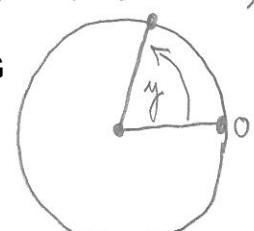
$$\varphi'(0) = \lim_{t \rightarrow 0} \frac{\varphi(t) - \varphi(0)}{t} \stackrel{E}{=} \lim_{t \rightarrow 0} E \left(\frac{e^{i \cdot t \cdot X} - 1}{t} \right) = \textcircled{*}$$

NOTE:
$$\left| \frac{e^{i \cdot t \cdot X} - 1}{t} \right| \stackrel{F}{\leq} |X|$$
, BECAUSE

$|e^{iy} - e^{i \cdot 0}| \leq |y|$ (CHORD IS SHORTER THAN ARCLENGTH)

$$\textcircled{*} \underset{H}{=} E \left(\lim_{t \rightarrow 0} \frac{e^{i \cdot t \cdot X} - 1}{t} \right) \stackrel{G}{=} E(i \cdot X)$$

DOM. CONV. THM., $E(|X|) < +\infty$ PAGE 30



THM: (CHAR. F'N. CHARACTERIZES DISTRIBUTION):

IF $\boxed{\varphi_X(t) \equiv \varphi_Y(t)}$ THEN

A

↑
 $\forall t \in \mathbb{R}$

$\boxed{P(X \leq x) \equiv P(Y \leq x)}$

B

↑
 $\forall x \in \mathbb{R}$

PROOF: LATER.

THM: $\boxed{X_m \Rightarrow Y} \Leftrightarrow \boxed{\forall t \in \mathbb{R}: \varphi_{X_m}(t) \rightarrow \varphi_Y(t)}$

C D

PROOF: \Rightarrow : IF $t \in \mathbb{R}$, LET $g(x) = e^{i \cdot t \cdot x}$

THEN $\varphi_{X_m}(t) = E(g(X_m)) \xrightarrow[E \rightarrow \infty]{} E(g(Y)) = \varphi_Y(t)$

(SEE THM ON PAGE 81, g IS BOUNDED, CONTINUOUS)

\Leftarrow : LATER.

F APPLICATION: IF $X_m \sim \text{BIN}(n, \frac{\lambda}{n})$ THEN

$X_m \Rightarrow Y$, WHERE $Y \sim \text{POI}(\lambda)$

PROOF: $X_m \underset{G}{=} Y_1 + \dots + Y_n$, WHERE

Y_1, \dots, Y_n ARE I.I.D., $\text{BER}\left(\frac{\lambda}{n}\right)$, THUS

$E(e^{i \cdot Y_n}) \underset{H}{=} \frac{\lambda}{n} \cdot e^{i \cdot t \cdot 1} + \left(1 - \frac{\lambda}{n}\right) \cdot e^{i \cdot t \cdot 0}$

$$\Psi_{X_m}(t) \underset{\text{A}}{=} \bar{E} \left(e^{i.t.(Y_1^{(m)} + \dots + Y_m^{(m)})} \right) \underset{\text{B}}{=} \text{INDEP.}$$

$$E(e^{i.t.Y_1^{(m)}}) \cdot \dots \cdot E(e^{i.t.Y_m^{(m)}}) \underset{\text{C}}{=} \left(1 + \frac{\lambda}{m} \cdot (e^{it} - 1)\right)^m$$

$$\lim_{n \rightarrow \infty} \Psi_{X_n}(t) \underset{\text{D}}{=} \exp(\lambda \cdot (e^{it} - 1))$$

IF $Y \sim \text{POI}(\lambda)$, THEN

$$E(e^{i.t.Y}) \underset{\text{E}}{=} \sum_{k=0}^{\infty} e^{i.t.k} \cdot \frac{e^{-\lambda} \cdot \lambda^k}{k!} \underset{\text{F}}{=} e^{-\lambda} \cdot \sum_{k=0}^{\infty} \frac{(e^{it} \cdot \lambda)^k}{k!} \underset{\text{G}}{=}$$

$$= \exp(-\lambda + e^{it} \cdot \lambda) = \exp(\lambda \cdot (e^{it} - 1))$$

THUS $\boxed{\Psi_{X_n}(t) \xrightarrow[n]{\infty} \Psi_Y} \underset{\text{H}}{\Rightarrow} \boxed{X_n \Rightarrow Y} \checkmark$

ANOTHER APPLICATION:

THM: (WEAK LAW OF LARGE NUMBERS):

IF X_1, X_2, \dots I.I.D., $E(|X_2|) < +\infty$

$$S_m = X_1 + \dots + X_m \text{ THEN}$$

$$\boxed{\frac{S_m}{m} \underset{\text{J}}{\Rightarrow} m}$$

$$\text{WHERE } m := E(X_2) \underset{\text{K}}{=}$$



PROOF: $\varphi(t) = \mathbb{E}(e^{itX_2})$

φ IS DIFFERENTIABLE \Rightarrow THUS

$$\varphi(t) = \underset{A}{\varphi(0)} + t \cdot \varphi'(0) + \bar{o}(t)$$

AN ERROR TERM
THAT CONVERGES
TO ZERO FASTER
THAN t

\hookrightarrow

1 \Rightarrow

\hookrightarrow

i.m \Rightarrow

\curvearrowright

\Rightarrow

$$\varphi_{S_m}(t) = \underset{B}{\left(\varphi(t)\right)^m}$$

$$\varphi_{\frac{S_m}{n}}(t) = \underset{C}{\left(\varphi\left(\frac{t}{n}\right)\right)^m} = \underset{D}{=}$$

$$= \left(1 + \frac{t}{n} \cdot i.m + \bar{o}\left(\frac{t}{n}\right)\right)^m \underset{E}{\xrightarrow[n]{\downarrow \infty}} e^{i.t.m}$$

\hookrightarrow THIS IS THE CHAR. FUNCTION OF THE
RANDOM VARIABLE CONCENTRATED
ON THE NUMBER m , THUS

$$\left[\frac{S_m}{n} \underset{F}{\Rightarrow} m \right] \Rightarrow \checkmark$$