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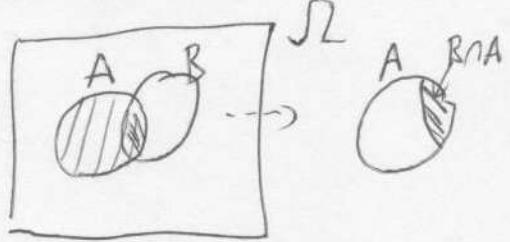
Conditional expectation

WARNING: This is HARD.

Introduction

Def: If $A, B \subset \Omega$ are events, $P(A) > 0$, then the conditional probability of B under the condition A is

$$P(B|A) = \frac{P(A \text{ and } B)}{P(A)}$$



[Philosophically: $\frac{\# \text{ favourable outcomes}}{\# \text{ all outcomes}}$]

Problem: What if $P(A) = 0$?

[Remark: A good question is why we are interested in that.
We will come to that later.]

Bad news 1: If $P(A) = 0$, then $P(B|A) = \frac{P(AB)}{P(A)}$ formally

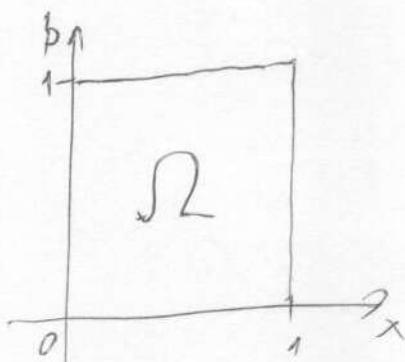
makes no sense: division by zero.

Bad news 2: If $P(A) = 0$, then $P(B|A)$ REALLY
makes no sense: not even intuitively!

Sampling example:

Consider the prob space $\Omega = [0, 1]^2$ with $P = \text{Leb}$,

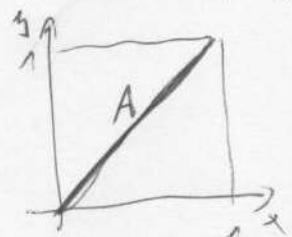
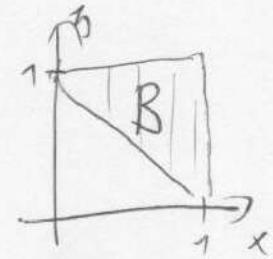
so $\omega \in (x, y) \in \Omega$ is uniformly chosen from $[0, 1]^2$



Now consider the events

$$B := \{(x, y) \in \Omega \mid x + y \geq 1\}.$$

$$A := \{(x, y) \in \Omega \mid y = x\}$$



How can we give sense to $P(B|A)$?

Idea 0: Approximate A with some A_ε such that P

$P(A_\varepsilon) > 0$ (although small), so $P(B|A_\varepsilon)$ makes sense,

and $\approx P(B|A)$ at least intuitively.

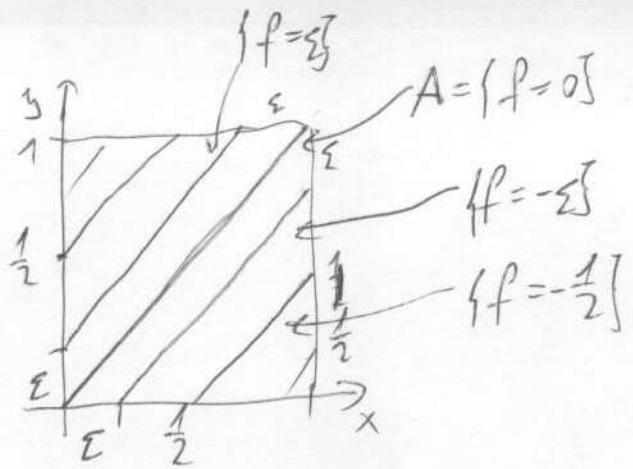
Q: How can we reasonably construct such an A_ε
in a natural way?

Spoiler: I will give two equally natural constructions
leading to completely different results.

Idea 1: Let $f: \Omega \rightarrow \mathbb{R}$, $f(x, y) = y - x$.

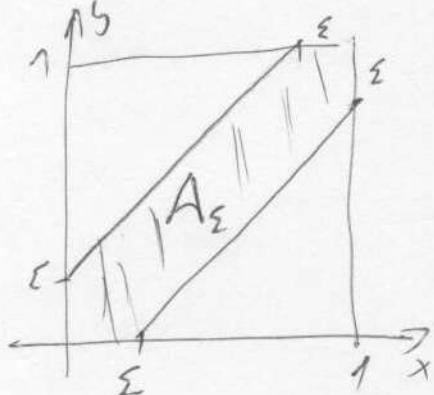
So $A = \{(x, y) \in \Omega \mid f(x, y) = 0\}$ is a level set of f .

Other level sets are also line segments!

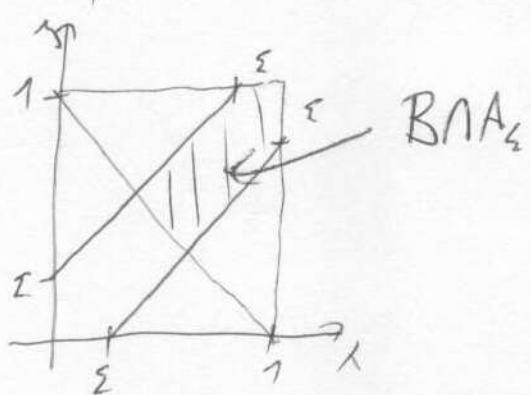


So let ~~A := \{(x,y) \in \mathbb{R}^2 \mid -\varepsilon \leq f(x,y) \leq \varepsilon\}~~ 3

$$A_\varepsilon := \{(x,y) \in \mathbb{R}^2 \mid -\varepsilon \leq f(x,y) \leq \varepsilon\}.$$



Easy to see:



$$P(B|A_\varepsilon) = \frac{1}{2} \text{ for } \forall \varepsilon > 0,$$

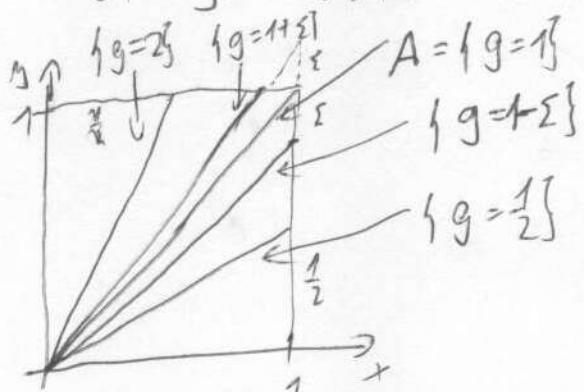
so

$$\underline{P}(B|A) := \lim_{\varepsilon \rightarrow 0} P(B|A_\varepsilon) = \frac{1}{2}.$$

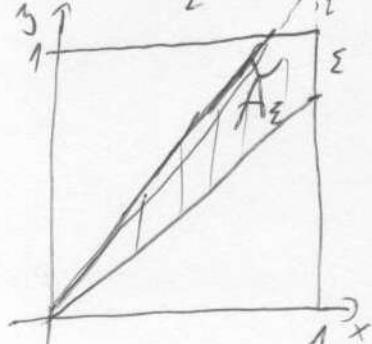
Def 2: Let $g: \mathbb{R}^2 \rightarrow \mathbb{R}$, $g(x,y) := \frac{y}{x}$ (defined \mathbb{R} -a.e.)

So $A = \{(x,y) \in \mathbb{R}^2 \mid g(x,y) = 1\}$ is again a level set

of g . Other level sets are also line segments:

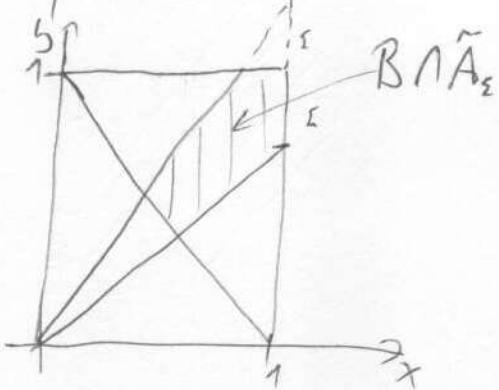


So let $\tilde{A}_\varepsilon := \{(x,y) \in \mathbb{R}^2 \mid 1 - \varepsilon \leq g(x,y) \leq 1 + \varepsilon\}$



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Easy to see:



$$P(B|A̅_\varepsilon) \approx \frac{3}{4}, \text{ so}$$

$\approx \text{triangle}$
 cut in half

$$\tilde{P}(B|A) := \lim_{\varepsilon \rightarrow 0} P(B|A̅_\varepsilon) = \underline{\underline{\frac{3}{4}}}.$$

Observation: The descriptions $\{f=0\} = A = \{g=1\}$

can be equally natural, depending on the context
 - whether A showed up while you studied $y-x$,
 or while you studied y/x .

~~Why we want the~~

Lesson to learn: For a single event $A \subset \mathbb{R}$ with $P(A)=0$,
 $P(B|A)$ can not be made sensible.

So we will have to define $\tilde{P}(B|A_i^2)$ simultaneously
 for an entire family of events $A_i^2 \subset \mathbb{R}$
 with $P(A_i^2)=0$

$\left[\begin{array}{l} \text{Say, in the above examples: } A_1^2 := \{f=2\} \quad (z \in [-1,1]) \\ \text{or } A_2^2 := \{g=2\} \quad (z \in [0,6]) \end{array} \right]$

Why we want to define $P(B|A)$ if $P(A)=0$

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Q: What is conditional prob. good for?

A: It's always used to calculate total probabilities.
nearly

~~Thm (Theorem of total~~ COUNTABLY MANY

Def: The sequence of events $A_1, A_2, A_3, \dots \subset \Omega$ is called a partition if always exactly 1 of them occurs:



$$\bigcup_i A_i = \Omega \quad (\text{covering})$$

$$A_i \cap A_j = \emptyset \quad \forall i \neq j \quad (\text{pairwise disjoint}).$$

Thm (Theorem of total probability) (countable)

Let $A_1, A_2, A_3, \dots \subset \Omega$ be a partition, and $B \subset \Omega$ an event

Then $P(B) = \sum_i P(A_i)P(B|A_i)$.

T.T.P.

Remark Convention: $0 \cdot \text{undefined} := 0$, so

if $P(A_i) = 0$, then $P(A_i)P(B|A_i) \xrightarrow{\text{convention}} 0$.

This is natural for 2 reasons:

- 1.) $P(B|A_i)$ makes no sense, but if it would, then it would surely be $\in [0, 1]$ and $0 \cdot \text{anything in } [0, 1] = 0$.
- 2.) $P(A_i)P(B|A_i) \xrightarrow{\text{should be}} \underbrace{P(B \cap A_i)}_{\subset A_i} \leq P(A_i)$, so $= 0$ if $P(A_i) = 0$.

Conclusion 1

It seems that we don't need to define $P(B|A_i)$ when $P(A_i) = 0$: it will/would be multiplied by 0 anyway.

Another tempting excuse:

$P(B|A)$ means: ~~what would~~ what would be the chance of B occurring, if A would occur?

But who cares what would happen, if A would occur?
If $P(A) = 0$, then A never occurs!

Problem with these:

1) This Theorem/Convention/Conclusion only hold as the partition is countable!

Adding "more than countably many zeroes" is not so simple.

2) Zero probability events do occur all the time:

If X is a continuous random variable, then $P(\{X=x\}) = 0$, but one of the (continuum many) $\{X=x\}$ is sure to happen.

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Conclusion 2:

If the theorem is the only important thing,
about conditional probability, then

Why not turn it into a definition?

A
axiomatic

Indeed: Our def. on conditional probability will essentially
be a generalization of the T.T.P. for uncountable
partitions

Idea of the construction

[WARNING again: this is a HARD construction/definition,
with many ingenious ideas.]

Idea 1: For given $B \in \mathcal{S}_2$ and partition $\{A_i\}_{i \in I}$ ~~and~~

we want to define all $P(B|A_i)$ simultaneously,

such that the T.T.P. holds:

$$P(B) = \sum_{i \in I} P(A_i) P(B|A_i).$$

However, this does not characterize $\{P(B|A_i)\}_{i \in I}$,
even when I is finite:

the equation $P(B) = \sum_{i \in I} P(A_i) x_i$ has many

solutions: a single equation for many variables $\{x_i\}_{i \in I}$.

So consider the following easy corollary:

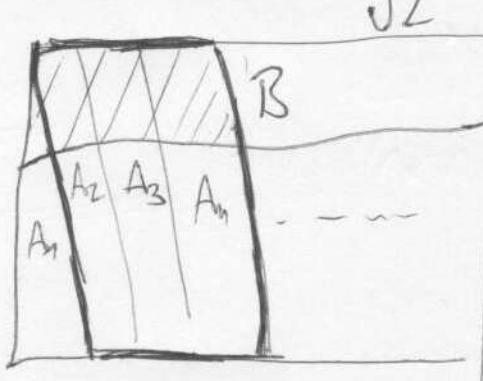
Thm (R.T.T.P. Extra): Let ~~A_1, A_2, \dots, A_n~~ $\{A_i\}_{i \in I}$ be a countable partition. Let $B \subset \Omega$ be an event.

Furthermore, let $C \subset \Omega$

be an event like this: →

$$C = \bigcup_{i \in J} A_i \text{ for some } J \subset I$$

or, equivalently,



$$C = A_2 \cup A_3 \cup A_4$$

for any $i \in I$ either $A_i \subset C$ or $A_i \cap C = \emptyset$.

$$\text{Then } P(B \cap C) = \sum_{i \in J} P(A_i) P(B|A_i)$$

[Weight of the striped  region above.]

This is what we will generalize.

Idea 2: Defining $P(B|A_i)$ simultaneously for all i
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means that the conditional probability is
not a number, but a function:

Cond. prob: ~~A_i~~

{partition elements} $\rightarrow \mathbb{R}[0, 1]$

$A_i \mapsto P(B|A_i)$

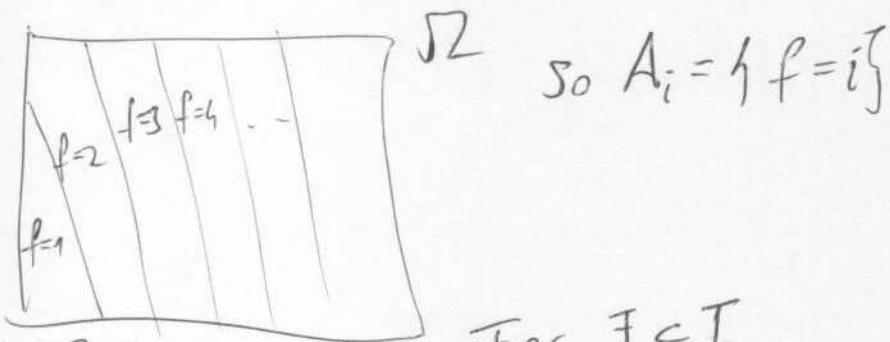
Q: How can we formulate such a function?

Answer 1: measure theorist's approach:

- Index your partition elements in your favourite way:
 Say, indexing with $i \in I$ will do, and then
 cond. prob. is a function $Y: I \rightarrow [0, 1]$
 $i \mapsto P(B|A_i)$

Equivalently: the A_i are level sets of the

function $f: \mathcal{S} \rightarrow I$ $f(w) = i$ for $w \in A_i$



Then the T.T.P. Extra reads: For $J \subset I$

$$P(B \cap \{f \in J\}) = \sum_{i \in J} P(A_i) Y(i).$$

This can then be generalized such that

$$\sum_{i \in J} P(A_i) Y(i) \text{ becomes } \int_J Y(i) d\gamma(i)$$

↓
some measure
on J .

This is doable.

This is a standard construction
in measure theory, called conditional measure.

This is not what we will do.

Disadvantage: The indexing is quite arbitrary:

- elements of the partition can be reordered
- equivalently: many functions $f: J \rightarrow$ ^{some} set I
produce the same level sets
- brings in the extra object I , which has no real meaning.

Answer 2: Probabilist's approach

[This is what we will do] [Slightly more general
than the previous]

Assigning numbers to partition elements $A_i \subset \Omega$

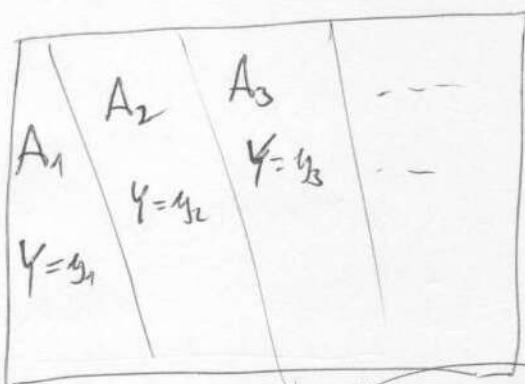
is equivalent to

$$Y: A_i \mapsto P(B|A_i)$$

is equivalent to ~~def~~ giving a function Y

- defined on Ω , so $Y: \Omega \rightarrow \{0, 1\}$

- but constant on each A_i : $Y|_{A_i} = Y_{y_i} := P(B|A_i)$



Now Thm T.T.P. Extra reads: (for $C = \bigcup_{i \in I} A_i$)

$$P(B \cap C) = \sum_{i \in I} P(A_i) y_i = \sum_{i \in I} \int_{A_i} Y(w) dP(w) = \int_C Y dP$$

\uparrow
 $y_i = y_j \text{ or } A_i \subset \Omega$

Moreover: **GREAT OBSERVATION:**

Let $\mathcal{G} = \sigma(\{A_i\}_{i \in I})$ be the σ -algebra generated by the partition

Then • "Y is constant on each A_i " means Y is measurable w.r.t. \mathcal{G} 12

• " $C = \bigcup_{i \in J} A_i$ for some $J \subset I$ " means $C \in \mathcal{G}$

So Thm T.T.P. Extra reads:

Thm (Theorem of Total probability Extra, equivalent 2nd version)

Let (Ω, \mathcal{F}, P) be a prob. space.

let $\{A_i\}_{i \in I}$ be a countable partition of Ω . (into events $A_i \in \mathcal{F}$)

let $\mathcal{G} = \sigma(\{A_i\}_{i \in I}) \subset \mathcal{F}$ be the sub- σ -algebra generated by the partition.

Let $B \in \mathcal{F}$ be any event.

Define $Y: \Omega \rightarrow [0, 1]$ as $Y(\omega) := P(B|A_i)$ when $\omega \in A_i$.

Notice: The good old $P(B|A_i) := \frac{P(A_i \cap B)}{P(A_i)}$ is well defined whenever

$P(A_i) > 0$. Since the partition is countable (by assumption), this $Y(\omega)$ is well defined for P -almost every ω .

Then • Y is \mathcal{G} -measurable, and

• for any $C \in \mathcal{G}$ $P(B \cap C) = \int_C Y dP$

This statement can be generalized to define the conditional probability for non-countable partitions.

This is ALMOST what we will do.

Last idea (quite obvious):

~~It's no extra diff~~ Without extra difficulty, we can immediately define the conditional probability expectation. Then the cond. prob. will be a special case:

$$P(B|A) = E(1_B|A).$$

Def: Let (Ω, \mathcal{F}, P) be a prob. space, $X: \Omega \rightarrow \mathbb{R}$ measurable, and assume that $E X$ exists. Let $A \in \mathcal{F}$, $P(A) > 0$.

Then the conditional expectation of X under the condition

A is $E(X|A) = \frac{\int_A X dP}{P(A)}$ [weighted average of values of X on A .]

Thm (Theorem of Total expectation) (T.T.E.)

Let (Ω, \mathcal{F}, P) be a prob. space, $X: \Omega \rightarrow \mathbb{R}$ measurable s.t. $E X$ exists.

Let $\{A_i\}_{i \in I}$ be a countable partition of Ω .

Then $E X = \sum_{i \in I} P(A_i) E(X|A_i)$

[with the convention $0 \cdot \text{undefined} := 0$]

Proof: trivial from the definition.

As before, a trivial (equivalent) generalization:

Thm (T.T.E Extra): Under the same assumptions,

if $C = \bigcup_{i \in I} A_i$ for some $I \subset I$, then

$$\int_C X dP = \sum_{i \in I} P(A_i) E(X|A_i).$$

And, just as before, $\sum_{i \in I} P(A_i) \cdot (\dots) \rightsquigarrow \int_C (\dots) dP$

Thm (Theorem of total expectation Extra, equivalent 2nd version)

Let (Ω, \mathcal{F}, P) be a prob. space.

Let $\{A_i\}_{i \in I}$ be a countable partition of Ω .

Let $\mathcal{G} = \sigma(\{A_i\}_{i \in I})$ be the generated σ -algebra.

Let $X: \Omega \rightarrow \mathbb{R}$ and assume that $E X$ exists.

Define $Y: \Omega \rightarrow [\alpha, \beta]$ as $Y(\omega) := E(X|A_i)$ when $\omega \in A_i$.

[Again: this is well-defined for P -a.e. ω .]

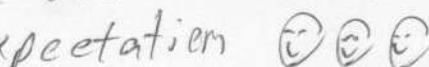
Then • Y is \mathcal{G} -measurable, and

$$\text{• for any } C \in \mathcal{G} \quad \int_C X dP = \int_C Y dP$$

Notice!! If $\int_C X dP = \int_C Y dP$ would hold for every $C \in \mathcal{F}$,

that would mean $Y = X$ ~~almost surely~~.

But we only claim it for $C \in \mathcal{G}$!!

We are (finally) ready to define the conditional expectation 

Def: Konditional expectation with respect to σ
 σ -algebra)

Let (Ω, \mathcal{F}, P) be a prob. space, $\mathcal{G} \subset \mathcal{F}$ a sub- σ -algebra. Let $X: \Omega \rightarrow \mathbb{R}$ integrable.

The function $Y: \Omega \rightarrow \mathbb{R}$ is called the conditional expectation of X w.r.t. \mathcal{G} , if

- 1) Y is \mathcal{G} -measurable, and
- 2) for any $A \in \mathcal{G}$ $\int_X dP = \int_A Y dP$.

Notation: $Y = E(X|\mathcal{G})$

Remark: Instead of „the conditional expectation“ it would be fair to say „a version of the conditional expectation“, since we haven't (yet) checked uniqueness.

Remark': If $Y = Z$ a.s., then $\int_A Y dP = \int_A Z dP$, so $E(X|\mathcal{G})$ can only be unique up to modification on 0-measure sets (respecting \mathcal{G} -measurability)

Example:

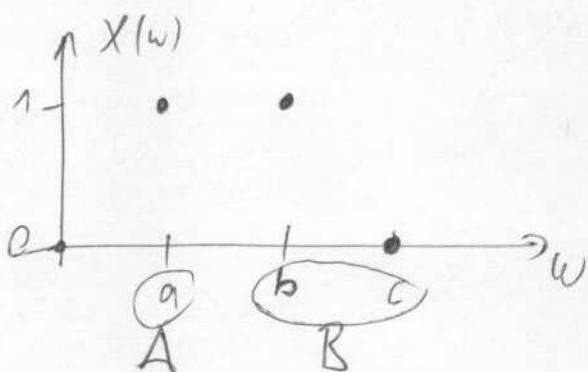
Let $\Omega = \{a, b, c\}$, $X = \mathbb{1}_{\{a, b\}}$, so $X(a) = 1$

$P = \frac{1}{3} X$: uniform prob.

$$P(\{a\}) = P(\{b\}) = P(\{c\}) = \frac{1}{3}$$

$\mathcal{F} = 2^{\Omega}$, discrete σ -algebra.

$\mathcal{G} = \{\emptyset, \{a\}, \{b, c\}, \Omega\}$: this is generated by
the partition $\{\{a\}, \{b, c\}\}$



$$\text{Set } A = \{a\}$$

$$B = \{b, c\}$$

Then $Y = \Omega \rightarrow \mathbb{R}$ defined as

$$Y|_A := 1 \quad Y|_B := \frac{1}{2} \quad \text{so} \quad Y(a) = 1 \quad Y(b) = \frac{1}{2} \quad Y(c) = \frac{1}{2}$$

will do for $Y = E(X|\mathcal{G})$:

$$\int_{\emptyset} X dP = 0 = \int_{\emptyset} Y dP \quad \checkmark$$

$$\int_A X dP = P(\{a\}) X(a) = \frac{1}{3} \cdot 1 = P(\{a\}) Y(a) = \int_A Y dP \quad \checkmark$$

$$\int_B X dP = \frac{1}{3} \cdot 1 + \frac{1}{3} \cdot 0 ; \quad \int_B Y dP = \frac{1}{3} \cdot \frac{1}{2} + \frac{1}{3} \cdot \frac{1}{2} \quad \text{EQUAL} \quad \checkmark$$

$$\int_{\Omega} X dP = \dots = \frac{2}{3} = \int_{\Omega} Y dP \quad \checkmark$$

Of course: we chose

$$\begin{aligned} Y|_A &:= \text{average of } X \text{ on } A = \{a\} \quad (\text{boring}) \quad \left. \begin{array}{l} \text{good old} \\ \text{conditional} \end{array} \right\} \\ Y|_B &:= \text{average of } Y \text{ on } B = \{b, c\} = \frac{1+0}{2}. \quad \left. \begin{array}{l} \text{expectation} \\ \text{version} \end{array} \right\} \end{aligned}$$

Thm (Theorem of total expectation extra, 3rd equivalent version)

Let (Ω, \mathcal{F}, P) be a prob. space.

Let $\{A_i\}_{i \in I}$ be a countable partition of Ω .

Let $G = \sigma(\{A_i\}_{i \in I}) \subset \mathcal{F}$ be the generated sub- σ -algebra.

Let $X: \Omega \rightarrow \mathbb{R}$ integrable

Define $Y: \Omega \rightarrow \mathbb{R}$ as $Y(\omega) := E(X|A_i)$ when $A_i \ni \omega \in A_i$

[Where $E(X|A_i)$ is the good old $E(X|A_i) := \frac{\int_{A_i} X dP}{P(A_i)}$]
 This is well defined a.s.

Then $E(X|G) = Y$.

So, indeed, $E(X|G)$ is the generalization of the old notion.

Very Special case: If $A \in \mathcal{F}$, $P(A) \neq 0$ and

$$G = \{\emptyset, A, A^c, \Omega\}, \text{ then } E(X|G)(\omega) = \begin{cases} E(X|A) & \text{if } \omega \in A \\ E(X|A^c) & \text{if } \omega \in A^c \end{cases}$$

Thm: The conditional expectation is unique up to changes on a set of measure 0:

If $Y: \mathcal{S} \rightarrow \mathbb{R}$ and $Z: \mathcal{S} \rightarrow \mathbb{R}$ are both versions of $E(X|G)$, then $P(Z=Y)=1$.

Proof: Y, Z are both G -measurable, so

$Y-Z$ is also G -measurable, so

$A := \{Y-Z > 0\} \in G$, so by definition of the conditional expectation

$$\int_A Y dP = \int_A X dP = \int_A Z dP \Rightarrow \int_A Y-Z dP = 0$$

$\nwarrow >0 \text{ on } A$

$$P(Y>Z) = P(A) = 0$$

$$\text{Similarly } P(Y<Z) = 0 \quad \square$$

Existence of $E(X|G)$

Thm: The conditional expectation exists:

For any $(\mathcal{S}, \mathcal{F}, P)$ prob-space, $G \subset \mathcal{F}$, $X: \mathcal{S} \rightarrow \mathbb{R}$ integrable

$\exists Y: \mathcal{S} \rightarrow \mathbb{R}$ such that $Y = E(X|G)$

$$\left[\begin{array}{l} \text{meaning } Y \in \mathcal{G} \text{ (G-measurable)} \\ \text{and } \int_A X dP = \int_A Y dP \quad \forall A \in \mathcal{G} \end{array} \right].$$

Proof: Assume 1st that $X \geq 0$.

Recall the Radon-Nikodym theorem: (with unusual notation)

~~If $(\mathcal{S}, \mathcal{G})$ is a meas~~

If $(\mathcal{S}, \mathcal{G})$ is a measurable space.

$\mu, \nu: \mathcal{G} \rightarrow \mathbb{R}$ σ -finite measures

and $\nu \ll \mu$,

then $\exists Y: \mathcal{S} \rightarrow \mathbb{R}^+$ \mathcal{G} -measurable

such that $\nu(A) = \int_A Y d\mu$ for all $A \in \mathcal{G}$.

Of course Y is \mathcal{G} -measurable;

Of course $A \in \mathcal{G}$:

there's no other σ -algebra in the assumptions.

But this measurability is of key importance for us now.

[The first time during the course]

So consider the measurable space $(\mathcal{S}, \mathcal{G})$ with the measures μ, ν on this space: $\mu, \nu: \underline{\mathcal{G}} \rightarrow \mathbb{R}^+$

$\mu(A) := P(A)$ for all $A \in \mathcal{G}$ - so $\mu = P|_{\mathcal{G}}$

$\nu(A) := \int_A X dP$ for $A \in \mathcal{G}$.

Lemma 1: These are both finite measures on $(\mathcal{S}, \mathcal{G})$.

Proof: easy, HW - use the def. of the measure
and basic properties of the integral.

In particular, $\mu(\mathcal{S}) = P(\mathcal{S}) = 1$

$$V(\mathcal{S}) = \int_{\mathcal{S}} X dP = \mathbb{E} X < \infty.$$

Lemma 2: $V \ll \mu$.

Proof: Obvious: If $\mu(A) = P(A) = 0$, then

$$V(A) = \int_A X dP = 0 \quad \checkmark$$

So applying the Radon-Nikodym thm gives:

$\exists Y: \mathcal{S} \rightarrow \mathbb{R}$ \mathcal{G} -measurable s.t. $\forall A \in \mathcal{G}$ $\$$

$$V(A) = \int_A X d\mu$$

$$\int_A X dP \stackrel{\text{def}}{=} \int_A Y d\mu$$

Lemma 3: For $A \in \mathcal{G} \subset \mathcal{F}$, $\mu = P|_{\mathcal{G}}$ $\int_A Y d\mu = \int_A Y dP$ ~~PROOF~~ \checkmark

integral on $(\mathcal{S}, \mathcal{G})$ integral on $(\mathcal{S}, \mathcal{F})$

Formally they are not the same, so we would need to check that they are equal — DON'T DO IT!

We are done when $X \geq 0$.

In the general case, write $X = X_+ - X_-$

positive part negative part,

so $X_+ \geq 0, X_- \geq 0$,

so ~~Y~~ $Y_+ := E(X_+ | \mathcal{G})$ and $Y_- := E(X_- | \mathcal{G})$ exist,

so $Y := Y_+ - Y_-$ will do (easy). \square

Def. (Conditional probability w.r.t. a σ -algebra)

Let (Ω, \mathcal{F}, P) be a prob. space, $A \in \mathcal{F}$, $\mathcal{G} \subset \mathcal{F}$ a sub-

σ -algebra. Then $P(A | \mathcal{G}) \stackrel{\text{def}}{=} E(\mathbb{1}_A | \mathcal{G})$

Properties of conditional expectation

1.) If ~~$\mathcal{G} = \emptyset$~~ , then $E(X | \mathcal{G}) = X$.

2.) More generally: If X is \mathcal{G} -measurable, then $E(X | \mathcal{G}) = X$

Proof: obvious - check the definition.

3.) If $\mathcal{G} = \{\emptyset, \Omega\}$ (the indiscrete σ -algebra), then

$E(X | \mathcal{G}) \equiv E(X)$ (constant function)

Proof: check the def.

4.) If $a, b \in \mathbb{R}$, X, Y are integrable then

$$\mathbb{E}(aX+bY|G) = a\mathbb{E}(X|G) + b\mathbb{E}(Y|G) \quad : \text{linearity}$$

Proof: check the def.

5.) Tower rule: Let $G_2 < G_1 < F$, so

G_2 is a smaller / rougher σ -algebra than G_1 .

Then $\mathbb{E}(\mathbb{E}(X|G_1)|G_2) = \mathbb{E}(X|G_2)$, No typo:

but also $\mathbb{E}(\mathbb{E}(X|G_2)|G_1) = \mathbb{E}(X|G_2)$ the smaller / rougher σ -algebra wins.

Phenomenon: ~~the~~ smaller σ -algebra = less information.