

Stochastic Differential Equations

Problem Set 1

Brownian Motion: Construction and Basic Properties

1.1 Let

$\varphi : \mathbb{R} \rightarrow \mathbb{R}_+$, $\varphi(x) := \frac{1}{\sqrt{2\pi}}e^{-x^2/2}$, be the standard normal density function,

$\Phi : \mathbb{R} \rightarrow [0, 1]$, $\Phi(x) := \int_{-\infty}^x \varphi(y)dy$, be the standard normal distribution function.

Prove that for any $x > 0$

$$\left(\frac{1}{x} - \frac{1}{x^3}\right)\varphi(x) < 1 - \Phi(x) < \frac{1}{x}\varphi(x).$$

Hint: Compare the derivatives.

1.2 For every $n \in \mathbb{N}$ let $X_1^{(n)}, X_2^{(n)}, \dots, X_n^{(n)}$ be i.i.d. normal random variables with

$$\mathbf{E}\left(X_j^{(n)}\right) = 0, \quad \mathbf{Var}\left(X_j^{(n)}\right) = \frac{1}{n}, \quad j = 1, \dots, n.$$

Define the stochastic process $t \mapsto B^{(n)}(t)$, $t \in [0, 1]$ as follows:

$$B^{(n)}(t) := \sum_{j=1}^{\lfloor nt \rfloor} X_j^{(n)}.$$

(a) Compute the expectations and covariances

$$\mathbf{E}\left(B^{(n)}(t)\right) = ?, \quad \mathbf{Cov}\left(B^{(n)}(t), B^{(n)}(s)\right) = ?, \quad s, t \in [0, 1],$$

and their limits as $n \rightarrow \infty$.

(b) What is the joint distribution of the random variables $\{B^{(n)}(t) : t \in [0, 1]\}$?

(c) Let

$$\delta_n := \max \{ |B^{(n)}(t+) - B^{(n)}(t-)| : t \in [0, 1] \}.$$

(In plain words: δ_n is the largest jump discontinuity of the process $\{B^{(n)}(t) : t \in [0, 1]\}$.)

Prove that for any fixed $\varepsilon > 0$,

$$\lim_{n \rightarrow \infty} \mathbf{P}(\delta_n \geq \varepsilon) = 0.$$

Hint: Note that $\delta_n = \max_{1 \leq j \leq n} |X_j^{(n)}|$ and use the upper bound from problem 1.1.

1.3 For every $n \in \mathbb{N}$ let $Y_1^{(n)}, Y_2^{(n)}, \dots, Y_n^{(n)}$ be i.i.d. Poisson random variables with parameter $1/n$. So,

$$\mathbf{E}(Y_j^{(n)}) = \frac{1}{n}, \quad \mathbf{Var}(Y_j^{(n)}) = \frac{1}{n}, \quad j = 1, \dots, n.$$

Define the stochastic process $t \mapsto B^{(n)}(t)$, $t \in [0, 1]$ as follows:

$$Z^{(n)}(t) := \sum_{j=1}^{\lfloor nt \rfloor} \left(Y_j^{(n)} - \frac{1}{n} \right).$$

(a) Compute the expectations and covariances

$$\mathbf{E}(Z^{(n)}(t)) = ?, \quad \mathbf{Cov}(Z^{(n)}(t), Z^{(n)}(s)) = ?, \quad s, t \in [0, 1],$$

and their limits as $n \rightarrow \infty$.

(b) What is the joint distribution of the random variables $\{Z^{(n)}(t) : t \in [0, 1]\}$? Explain in plain words.

(c) Let

$$\delta_n := \max \{ |Z^{(n)}(t+) - Z^{(n)}(t-)| : t \in [0, 1] \}.$$

(In plain words: δ_n is the largest jump discontinuity of the process $\{Z^{(n)}(t) : t \in [0, 1]\}$.)

Compute, for $\varepsilon > 0$ fixed,

$$\lim_{n \rightarrow \infty} \mathbf{P}(\delta_n \geq \varepsilon).$$

Hint: Note that $\delta_n = \max_{1 \leq j \leq n} |Y_j^{(n)}|$ and use all you know about Poisson random variables.

1.4 Interpret the results of problems 1.2, respectively, 1.3.

1.5 (a) Let Y_1, Y_2, \dots, Y_n be random variables with $\mathbf{E}(Y_j) = 0$ and $\mathbf{Cov}(Y_i, Y_j) =: c_{i,j}$. Assume that the covariance matrix $C := (c_{i,j})_{i,j=1}^n$ is non-degenerate, $\det(C) \neq 0$. Prove that the random variables Y_1, Y_2, \dots, Y_n are *jointly Gaussian* if and only if there exist i.i.d. $\mathcal{N}(0, 1)$ -distributed random variables X_1, X_2, \dots, X_n and real coefficients $(a_{i,j})_{i,j=1}^n$ such that

$$Y_i = \sum_{j=1}^n a_{ij} X_j.$$

Hint: Express the matrix $A = (a_{i,j})_{i,j=1}^n$ from the covariance matrix $C = (c_{i,j})_{i,j=1}^n$.

(b) Let $t \mapsto B(t)$ be standard 1d Brownian motion and $0 \leq t_1 \leq t_2 \leq \dots \leq t_n$. Explain why it follows from the definition of Brownian motion (i.e. independent and Gaussian increments) that the random variables $B(t_1), B(t_2), \dots, B(t_n)$ have jointly Gaussian distribution.

(c) Determine the covariance matrix of the random variables $B(t_1), B(t_2), \dots, B(t_n)$.

1.6 Let $t \mapsto B(t)$ be standard 1d Brownian motion. Prove that the following processes are also standard 1d Brownian motions:

(a) The rescaled process: $X(t) := a^{-1/2}B(at)$, where $a > 0$ is fixed parameter.

(b) The time reversed process: $Y(t) := tB(1/t)$.

(c) The backwards process: $Z(t) := B(T) - B(T - t)$, where $T > 0$ is fixed and $t \in [0, T]$.

Hint: Prove that the processes $X(t), Y(t), Z(t)$ are Gaussian and compute their covariances.

1.7 For $j = 1, \dots, n$, let $t \mapsto B_j(t)$, be independent 1d Brownian motions with variance σ_j^2 , and a_j fixed real numbers. Prove that the process $t \mapsto Z(t) := \sum_{j=1}^n a_j B_j(t)$ is also a 1d Brownian motion. Determine the variance of the process $Z(t)$.

1.8 Let $t \mapsto B(t)$ be standard 1d Brownian motion. Determine (without painful computations) the conditional probability

$$\mathbf{P}(B(2) > 0 \mid B(1) > 0).$$

1.9 Show that 1d Brownian motion changes sign infinitely many times in any time interval $[0, \delta]$ of positive length δ .

1.10 The *Brownian meander* process.

(a) Let $\varepsilon > 0$ be fixed. Using the *reflection principle* prove that for any $x > 0, t > 0$

$$\mathbf{P}(B(t) \geq x - \varepsilon \mid \min_{0 \leq s \leq t} B(s) \geq -\varepsilon) = \frac{\Phi((-x + \varepsilon)/\sqrt{t}) - \Phi((-x - \varepsilon)/\sqrt{t})}{2\Phi(\varepsilon/\sqrt{t}) - 1}. \quad (*)$$

(b) Letting $\varepsilon \rightarrow 0$ in the previous formula prove that the conditional density of $B(t)$, given $\{B(s) \geq 0 : s \in [0, t]\}$ is

$$\frac{x}{t} \exp\{-x^2/(2t)\} \mathbb{1}_{\{x>0\}}.$$

Remark: Note that the probability of the condition is zero (see problem 1.9). So, strictly speaking, the conditional distribution doesn't make sense. Indeed, let us *define* the conditional probabilities as

$$\mathbf{P}(B(t) \in A \mid B(s) \geq 0 : s \in [0, t]) := \lim_{\varepsilon \searrow 0} \mathbf{P}(B(t) \in A \mid \min_{0 \leq s \leq t} B(s) \geq -\varepsilon).$$

Brownian motion conditioned to stay positive in this sense is called *Brownian meander*.

1.11 On the Hilbert space $\mathcal{L}^2([0, 1], dx)$ define the self-adjoint compact (actually: Hilbert-Schmidt) operator

$$Kf(t) := \int_0^1 \min\{t, s\} f(s) ds.$$

Prove that

$$\lambda_n = \frac{4}{\pi^2(2n-1)^2}, \quad \psi_n(t) = \sqrt{2} \sin\left(\frac{\pi(2n-1)}{2}t\right), \quad n = 1, 2, \dots$$

are eigenvalues and eigenvectors of the operator K .

1.12 Let ξ be a standard normal random variable and define, for $\lambda < 1$

$$\psi(\lambda) := \log \mathbf{E}(\exp\{\lambda(\xi^2 - 1)/2\}).$$

Prove that

$$\psi(\lambda) = -\frac{1}{2}(\log(1-\lambda) + \lambda),$$

and investigate the analytic properties of the function $\psi(\cdot)$ (convexity, minima, asymptotes, ...). Plot the graph of the function $\lambda \mapsto \psi(\lambda)$.

1.13 Show that the function

$$\phi : \mathbb{R}_+ \times \mathbb{R} \rightarrow \mathbb{R}, \quad \phi(t, x) := \frac{1}{\sqrt{t}} \varphi\left(\frac{x}{\sqrt{t}}\right)$$

solves the heat equation

$$\partial_t \phi(t, x) = \frac{1}{2} \partial_x^2 \phi(t, x).$$

1.14 Exercise 1 implies that if ξ is a standard Gaussian random variable and $x \geq 1$, then

$$\mathbf{P}(|X| \geq x) \leq \sqrt{\frac{2}{\pi}} e^{-\frac{x^2}{2}}.$$

Use this to show that if ξ_1, ξ_2, \dots are i.i.d. standard Gaussian, then, with probability 1, the event $\{|\xi_n| > 2 \ln n\}$ occurs for at most finitely many n -s.

1.15 *Paul Lévy construction of the Wiener process. Second version: wrong constants corrected – sorry.* In a possible construction of the Wiener process (or Brownian motion) on $[0, 1]$ we define a sequence of piecewise linear continuous random functions so that we first define f_n at dyadic rationals that are multiples of $\frac{1}{2^n}$, inheriting every second value (at multiples of $\frac{1}{2^{n-1}}$) from f_{n-1} , and setting the values at the remaining points (of the form $\frac{2k-1}{2^n}$) to be the average of the two neighbouring values, plus an independent Gaussian random value with mean 0 and variance $\frac{1}{2^{n+1}}$. Then we extend f_n to $[0, 1]$ piecewise linearly.

Formally: we take independent standard Gaussian random variables ξ_0 and $\xi_{n,k}$ where $n = 1, 2, \dots$ and $k = 1, 2, \dots, 2^{n-1}$. Then

- In the 0th step we fix $f_0(0) = 0$ and $f_0(1) = \xi_0$. We connect these two values linearly.
- In the 1st step we leave $f_1(0) = f_0(0)$ and $f_1(1) = f_0(1)$, but also set $f_1(\frac{1}{2}) = f_0(\frac{1}{2}) + \frac{1}{2}\xi_{1,1}$. We connect these three values linearly.
- ... in the n th step we leave $f_n(\frac{k}{2^{n-1}}) = f_{n-1}(\frac{k}{2^{n-1}})$ for $k = 0, 1, \dots, 2^{n-1}$, but also set $f_n(\frac{k-\frac{1}{2}}{2^{n-1}}) = f_{n-1}(\frac{k-\frac{1}{2}}{2^{n-1}}) + \frac{1}{\sqrt{2^{n+1}}}\xi_{n,k}$ for $k = 1, \dots, 2^{n-1}$. We connect these $2^n + 1$ values linearly.

Notice that, in this construction, the difference $g_n := f_{n+1} - f_n$ is the sum of 2^n “tent” maps with disjoint supports and i.i.d. Gaussian “heights”.

(a) Use the statement of Exercise 14 to show that, with probability 1, the series

$$\lim_{n \rightarrow \infty} f_n = f_0 + \sum_{n=0}^{\infty} g_n$$

is uniformly absolutely convergent.

(b) Check that the limit is a Wiener process.

1.16 (Based on Exercise 8.1.3. from [1].) Let $B(t)$ be a standard Brownian motion (Wiener process). Fix $t > 0$ and for $n = 0, 1, 2, \dots$ let

$$V_n = \sum_{m=0}^{2^n-1} \left(B\left(\frac{m+1}{2^n}t\right) - B\left(\frac{m}{2^n}t\right) \right)^2.$$

Calculate the expectation and the variance of V_n . Use the Borel-Cantelli lemma to show that $V_n \rightarrow t$ almost surely as $n \rightarrow \infty$.

References

- [1] Durrett, R. *Probability: Theory and Examples*. **4th** edition, Cambridge University Press (2010)